

Review of Ch 2-5

Refer to Exam 1 review, 1-12.

1) know Ch 2-3: $X \sim \text{Exponential}(\mu)$

$$f(x) = \mu e^{-\mu x}, \quad x > 0$$

$$E(X) = \frac{1}{\mu}, \quad V(X) = \frac{1}{\mu^2}$$

$$F(x) = 1 - e^{-\mu x}, \quad x > 0, \quad {}_x p_x = S(x) = e^{-\mu x}, \quad x > 0.$$

In this class, usually $\mu < 1, E(X) > 1$.2) know $X \sim U(\theta_1, \theta_2)$ especially the)De Moivre(θ) = $U(0, \theta)$ distribution.

$$f(x) = \frac{1}{\theta}, \quad 0 \leq x \leq \theta.$$

$$E(X) = \frac{\theta}{2}, \quad V(X) = \frac{\theta^2}{12}$$

$$F(x) = \frac{x}{\theta}, \quad 0 \leq x \leq \theta$$

3) Usually $x \geq 0$ is nonnegative and interest is in the survival function $S(x) = 1 - F(x)$.4) cdf $F(x) = P(X \leq x)$, $S(x) = P(X > x)$,pdf $f(x) = F'(x)$, force of mortality $\mu(x) =$ hazard rate function, $E(X) = \bar{e}_0$

5) (x) denotes a person alive at age x . (1.5)

6) Let $X = T_0$ where $T_x = T(x)$ is time until failure (death) for a person alive at age x . $T_0 = x + T_x$ given $T_0 > x$.

7)

Let $G_x = G_{T_x}$

G is T, S, F, μ or t . If there is no subscript x , then $G = G_0$. Let $t > 0$.

$$i) {}_tP_x = S_x(t) = \frac{S_0(x+t)}{S_0(x)} = \exp\left(-\int_x^{x+t} \mu_y dy\right) = \exp\left(-\int_0^t \mu_{x+w} dw\right)$$

$$ii) {}_tq_x = F_x(t) = 1 - {}_tP_x = 1 - \frac{S_0(x+t)}{S_0(x)} = P(T_x \leq t)$$

$$= P(T_0 \leq x+t \mid T_0 > x)$$

$$iii) {}_tP_x \mu_{x+t} = f_x(t) = \frac{f_0(x+t)}{S_0(x)} = \frac{d}{dt} F_x(t) = -\frac{d}{dt} S_x(t)$$

$$iv) \mu_{x+t} = \mu_x(t) = \mu_0(x+t) = \frac{f_0(x+t)}{S_0(x+t)} = \frac{f_x(t)}{S_x(t)}$$

* These quantities are nonnegative ≥ 0 .

LHS = actuarial notation

8)* If $t=1$, the subscript is

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usually suppressed. So $p_x = {}_1p_x$ and $q_x = {}_1q_x$.

9) The complete expectation of life at age x
= expected future lifetime at age x =

$$\overset{\circ}{e}_x = E(T_x) = \int_0^{\infty} t f_x(t) dt = \int_0^{\infty} t p_x dt = \int_0^{\infty} S_x(t) dt.$$

Note that $\overset{\circ}{e}_0 = E(T_0)$.

10) $V(T_x) = E(T_x^2) - [E(T_x)]^2$ where

$$E(T_x^2) = \int_0^{\infty} t^2 f_x(t) dt = 2 \int_0^{\infty} t t p_x dt \\ = 2 \int_0^{\infty} t S_x(t) dt. \quad \text{Use } x=0 \text{ to find } E(T_0^2).$$

11) If $T_0 \sim U(0, w)$, then T_x has a
De Moivre $(w-x)$ distribution; $T_x \sim U(0, w-x)$,
with support $0 < t < w-x$. For such a t ,

$$S_x(t) = t p_x = \frac{w-x-t}{w-x} = 1 - \frac{t}{w-x}$$

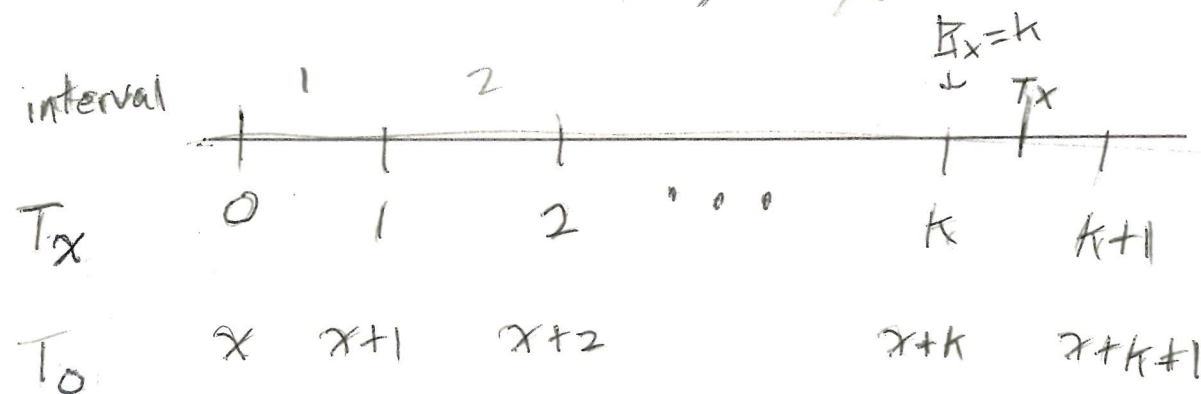
$$\mu_x(t) = \mu_{x+t} = \frac{1}{w-x-t} \quad E(T_x) = \overset{\circ}{e}_x = \frac{w-x}{2}.$$

Often need to recognize this distribution
from $t p_x$.

12] The curtate duration at failure RV (2.5)

$$K_x = \lfloor T_x \rfloor \quad \text{where } \lfloor 7.8 \rfloor = 7.$$

$K_x = k$ means person with T_x died in the time interval $(k, k+1)$.



at time $t=40$,
if $x=40$
 T_0 is 40
but $T_{40} = 0$

So person with T_0 died in time interval $(x+k, x+k+1)$, given $T_0 > x$.

13] The pmf of K_x is $P(K_x = k) = P_{K_x}(k) =$
 $k|q_x = P(k < T_x < k+1) = P(x+k < T_0 < x+k+1 | T_0 > x)$
 $= kP_x - (k+1)P_x = F_x(k+1) - F_x(k) = S_x(k) - S_x(k+1).$

The curtate expectation of life

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at age x is $e_x = E[K_x]$.

$$E(K_x) = e_x \approx e_x + 0.5.$$

let $w = x|y = a$
so $T_x = T_0 | T_0 > x$

14) The probability that (x) will die between $x+n$ and $x+n+m$ is

$${}_n m q_x = P(x+n < T_0 \leq x+n+m | T_0 > x)$$

$$= {}_n p_x - {}_{n+m} p_x = {}_{n+m} q_x - {}_n q_x = {}_n p_x \cdot m q_{x+n}$$

For $m=1$, ${}_n | 1 q_x = {}_n | q_x = P(K_x = n)$

$$= P(x+n < T_0 < x+n+1 | T_0 > x).$$

15) multiplication rule: ${}_{n+m} p_x = {}_n p_x \cdot m p_{x+n}$

← "angle n"

16) ${}^o e_{x:\overline{n}|}$ = expected number of years

lived in $(x, x+n]$ by a randomly selected survivor to age x , and

$${}^o e_{x:\overline{n}|} = \int_0^n {}_t p_x dt = \int_0^n s_x(t) dt$$

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17) i) The compound interest factor

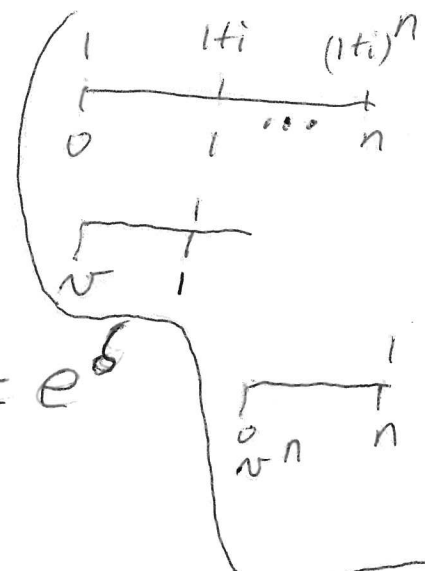
$$v = \frac{1}{1+i} \quad \text{and} \quad 0 < v < 1.$$

ii) The effective rate of interest

$$i = \frac{1-v}{v} > 0.$$

iii) The force of interest

$$\delta = \log(1+i) > 0, \quad 1+i = e^\delta$$



know: so $v = e^{-\delta}$.

iv) The effective rate of discount $d = \frac{i}{1+i} = iv = 1-v > 0.$

18) A life insurance model has benefit function b_x and a discount function v_x where x = length of time from issue of insurance until death.

Often $v_x = v^x$ and $b_x = 1$ unit.

The present value function $z_x = b_x v_x$.

The present value RV $Z = z_{T_x}$
or $Z = z_{K_x+1}$. Often let $T = T_x$.

19) $E(Z) =$ actuarial present value (APV)
 $=$ expected present value (EPV) =
 net single premium (NSP).

Z ,
 A For
 Insurance

20) Formulas are usually given for
 unit payment. Let $A = E(Z)$ and
 ${}^2A = E(Z^2)$. For nonunit payment C ,
 multiply the unit payment formula for A
 by C and the unit payment formula
 for 2A by C^2 .

21) Discrete insurance

i) whole life $Z_t = v^t, t \geq 0$

$$A_x = E(Z_x) = E(v^{K_x+1})$$

$${}^2A_x = E(Z_x^2)$$

$$\begin{cases} 1 & t \leq n \\ 0 & \text{else} \end{cases}$$

ii) n year term $Z_t = v^t I(t \leq n)$

$$Z_{x:\overline{n}|} = v^{K_x+1} I(K_x \leq n)$$

decorations

$$A'_{x:\overline{n}|} = E(Z'_{x:\overline{n}|})$$

$${}^2 A'_{x:\overline{n}|} = E[(Z'_{x:\overline{n}|})^2]$$

iii) n year deferred

$$Z_t = v^t I(t > n)$$

$$n|Z_x = v^{K_x+1} I(K_x \geq n)$$

$$n|A_x = E(n|Z_x)$$

$${}^2 n|A_x = E[(n|Z_x)^2]$$

iv) n year endowment

$$Z_t = \begin{cases} v^t & t \leq n \\ v^n & t > n \end{cases} = v^{\min(t, n)}$$

$$Z_{x:\overline{n}|} = \begin{cases} v^{K_x+1}, & K_x < n \\ v^n, & K_x \geq n \end{cases}$$

$$A_{x:\overline{n}|} = E(Z_{x:\overline{n}|}), \quad {}^2 A_{x:\overline{n}|} = E[(Z_{x:\overline{n}|})^2]$$

xv) discrete = continuous pure endowment insurance

$$Z_t = v^n I(t > n)$$

$$\ddot{Z}_{x:\overline{n}|} = \frac{v^n}{e^{-\delta n}} I(T_x > n)$$

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$$* A_{x:\overline{n}|} = E(\ddot{Z}_{x:\overline{n}|}) = e^{-\delta n} S_x(n) = e^{-\delta n} {}_n p_x$$

easy to compute

$${}^2 A_{x:\overline{n}|} = E[(\ddot{Z}_{x:\overline{n}|})^2] = e^{-2\delta n} S_x(n) = e^{-2\delta n} {}_n p_x$$

$$V(\ddot{Z}_{x:\overline{n}|}) = e^{-2\delta n} {}_n p_x {}_n q_x = e^{-2\delta n} S_x(n) F_x(n)$$

22) $\ddot{Z}_x = \ddot{Z}'_{x:\overline{n}|} + n|\ddot{Z}_x$, $A_x = A'_{x:\overline{n}|} + n|A_x$

whole term deferred

$$(\ddot{Z}_x)^2 = (\ddot{Z}'_{x:\overline{n}|})^2 + (n|\ddot{Z}_x)^2, \quad {}^2 A_x = {}^2 A'_{x:\overline{n}|} + {}^2 n|A_x$$

$\ddot{Z}_{x:\overline{n}|} = \ddot{Z}'_{x:\overline{n}|} + \ddot{Z}_{x:\overline{n}|}^1$, $A_{x:\overline{n}|} = A'_{x:\overline{n}|} + A_{x:\overline{n}|}^1$

n year endowment term pure endowment

$$(\ddot{Z}_{x:\overline{n}|})^2 = (\ddot{Z}'_{x:\overline{n}|})^2 + (\ddot{Z}_{x:\overline{n}|}^1)^2, \quad {}^2 A_{x:\overline{n}|} = {}^2 A'_{x:\overline{n}|} + {}^2 A_{x:\overline{n}|}^1$$

23) continuous insurance

note that 21 v) n year pure endowment insurance is both discrete and continuous insurance.

i) whole life $Z_t = v^t$, $t > 0$ (5.9)

$$\bar{Z}_x = v^T$$

$$T = T_x$$

$$\bar{A}_x = E(\bar{Z}_x) = E(e^{-\delta T}) = \int_0^{\infty} e^{-\delta t} f_x(t) dt = g(\delta)$$

$${}^2\bar{A}_x = E(\bar{Z}_x^2) = E(e^{-2\delta T}) = \int_0^{\infty} e^{-2\delta t} f_x(t) dt = g(2\delta)$$

ii) n year term $Z_t = v^t I(t \leq n)$

$$\bar{Z}'_{x:\overline{n}|} = v^T I(T \leq n)$$

bar
for
continuous

$$\bar{A}'_{x:\overline{n}|} = E(\bar{Z}'_{x:\overline{n}|}) = \int_0^n e^{-\delta t} f_x(t) dt$$

$${}^2\bar{A}'_{x:\overline{n}|} = E[(\bar{Z}'_{x:\overline{n}|})^2] = \int_0^n e^{-2\delta t} f_x(t) dt$$

iii) n year deferred $Z_t = v^t I(t > n)$

$${}_n|\bar{Z}_x = v^T I(T > n)$$

$${}_n|\bar{A}_x = E({}_n|\bar{Z}_x) = \int_n^{\infty} e^{-\delta t} f_x(t) dt$$

$${}^2{}_n|\bar{A}_x = E[({}_n|\bar{Z}_x)^2] = \int_n^{\infty} e^{-2\delta t} f_x(t) dt$$

(v) n year endowment $z_t = \begin{cases} v^t & t \leq n \\ v^n & t > n \end{cases} = v^{\min(t, n)}$ M402 6

$$\bar{Z}_{x:\overline{n}|} = \begin{cases} v^T & T \leq n \\ v^n & T > n \end{cases}$$

$$\bar{Z}_{x:\overline{n}|} = \bar{Z}'_{x:\overline{n}|} + Z_{x:\overline{n}|}^1$$

$$\bar{A}_{x:\overline{n}|} = \bar{A}'_{x:\overline{n}|} + A_{x:\overline{n}|}^1$$

← pure endowment does not get a bar since it is discrete and continuous

$$(\bar{Z}_{x:\overline{n}|})^2 = (\bar{Z}'_{x:\overline{n}|})^2 + (Z_{x:\overline{n}|}^1)^2$$

$${}^2\bar{A}_{x:\overline{n}|} = {}^2\bar{A}'_{x:\overline{n}|} + {}^2A_{x:\overline{n}|}^1$$

24)* If $T = T_x \sim \text{Exponential}(\mu)$,

then $\int_0^\infty \mu e^{-\mu t} dt = 1$ so $\int_0^\infty e^{-\mu t} dt = \frac{1}{\mu}, \mu > 0$.

If $b_t = e^{\theta t}$ and $v_t = v^t$, then $z_t = e^{\theta t} e^{-\delta t}$

and $\bar{Z} = z_T b_T$ and $\bar{A} = E(\bar{Z}) = E(e^{\theta T} e^{-\delta T})$

$$= \int_0^\infty e^{\theta t} e^{-\delta t} \mu e^{-\mu t} dt = \mu \int_0^\infty e^{-t(\mu + \delta - \theta)} dt$$

$$= \frac{\mu}{\mu + \delta - \theta} \quad \text{if } \mu + \delta - \theta > 0, \text{ and}$$

$$E[(\bar{Z})^j] = E \int_0^{\infty} (e^{\theta t} e^{-\delta t})^j \mu e^{-\mu t} dt =$$

6.5

$$\mu \int_0^{\infty} e^{-t} (t + \delta j - \theta j) dt = \frac{\mu}{\mu + \delta j - \theta j} \quad \text{if}$$

$\mu + \delta j - \theta j > 0$. Note $\theta = 0$ corresponds

* to unit payment: $\bar{A}_x \stackrel{E}{=} \frac{\mu}{\mu + \delta}$, ${}^2\bar{A}_x \stackrel{E}{=} \frac{\mu}{\mu + 2\delta}$

25] often use \int_0^n or \int_n^{∞} instead of \int_0^{∞}

ch 5 26] $a_{\overline{n}|} = \sum_{j=1}^n v^j = \frac{1-v^n}{i}$

$$\ddot{a}_{\overline{n}|} = \sum_{j=0}^{n-1} v^j = \frac{1-v^n}{d}$$

$$\bar{a}_{\overline{n}|} = \int_0^n v^t dt = \frac{1-v^n}{\delta}$$

a for annuity,
Y

27] i) A discrete immediate whole life annuity pays 1 unit at times 1, 2, ... as long as (x) survives. The present value RV

$$Y_x = a_{\overline{K_x}|} = \frac{1}{i} (1 - (1+i)^{-K_x})$$

$$APV = a_x = E Y_x, \quad V(Y_x) = \frac{{}^2A_x - (A_x)^2}{d^2}$$