

too long for MWF

E3 rev (102)

$0.5 \ln(95) = 2.2769$

1) You fit various loss models to 95 claim sizes using maximum likelihood. The fits maximizing the likelihood for a given number of parameters have the following log-likelihoods. Using the Schwarz Bayesian Criterion, what is the selected model? Hint: eliminate one of the 1 parameter models and one of the 2 parameter models by inspection.

closer to 0 is good

model	number of parameters	maximal loglikelihood
1	1	-487
2	1	-487.5
3	2	-487
4	2	-484.1
5	3	-482

model 1 is better than 2
model 4 is better than model 3
bad models

$\ln(Lr) - \frac{r}{2} \ln(95) =$

-489.2769

-489.770
 -491.9538

-488.6538 ←

-488.8307

model 4 with 2 parameters is selected

E3 rev (121)

2) Suppose that the full-credibility standard is defined such that aggregate claim dollars will be within 3% of its true value 95% of the time. The number of claims follows a Poisson distribution, and the claim severity distribution is exponential. Find the indicated full credibility standard in terms of number of claims, using limited fluctuation credibility.

$z_p = 1.96, k = .03, W = 5, e = nF$
 $CV(X) = 1$ (for EXP(θ))

$n_0 (1 + [CV(X)]^2) = 2 n_0$

$= 2 \left(\frac{1.96}{.03} \right)^2 = 8536.8889$

$= 2 (4268.4444)$

easy way!

table $v = \alpha\theta$ (142) (127) (132)
 $k = \frac{1}{\theta} = \frac{1}{1.2}$

Buhlmann credibility = Bayesian credibility

$E(X) = E[E(X|\lambda)] = E(\lambda) = 1.2$
 $\bar{x} = \frac{3+0}{2} = 1.5$ for this model

→ 3) The number of claims made by an individual insured in a year has a Poisson(λ) distribution where $\lambda \sim \text{gamma}(\alpha = 1, \theta = 1.2)$. Three claims are observed in year 1 and no claims are observed in year 2. Using Buhlmann credibility, estimate the number of claims in year 3. Hint: $n = 2, X_1 = 3, X_2 = 0$.

$\mu(\lambda) = E(X|\lambda) = \lambda = v\alpha = v(X|\lambda)$

$EPV = E[V(X|\lambda)] = E(\lambda) = E(\alpha\theta) = \alpha\theta = 1.2 = V = E(\lambda^2) - (E\lambda)^2$

$\alpha = \frac{VHM}{V} = \frac{V[E(X|\lambda)]}{V} = V(\lambda) = \alpha\theta^2 = 1.44$

$k = \frac{EPV}{VHM} = \frac{V}{\alpha} = \frac{1.2}{(1.2)^2} = \frac{1}{1.2} = 0.8333$
 $z = \frac{n}{n+k} = \frac{2}{2 + \frac{1}{1.2}} = 0.7059$

$P = z\bar{x} + (1-z)E(X) = EX|+z(x-E(X))$
 $= 1.2 + 0.7059(1.5 - 1.2) = \boxed{1.4118}$

E3 rev see 127

4) The table is used for a sample of 4 claim payments. Use the 4 step Kolmogorov Smirnov test for whether the claims are from an exponential(10) distribution.

E3 rev 96

$F(x) = 1 - e^{-x/10}$

x_i	$F_n(x_i)$	$F_n(x_i^-)$	$F^*(x_i)$	$\max(F_n(x_i) - F^*(x_i) , F_n(x_i^-) - F^*(x_i))$
2	1/4	0/4	0.181	0.181 = .181 - 0
5	2/4	1/4	0.393	0.143 = .393 - .25
13	3/4	2/4	.727	0.227 = .727 - .5 ← D
22	4/4	3/4	.889	0.139 = .889 - .75

i) H_0 EXP(10) is a good fit H_1 not H_0

ii) $D = 0.227$

iii) critical value = $\frac{k36}{\sqrt{n}} = \frac{1.36}{2} = 0.68$

$D < 0.68$, fail to reject H_0

iv) the EXP(10) dist is a good fit

(or not enough evidence to conclude EXP(10) is not a good fit)

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total claims k	number of policies n_k
0	5000
1	2100
2	750
3	100
4	50

$$\bar{X} = \hat{V} = \frac{0 + 2100 + 750(2) + 100(3) + 50(4)}{8000}$$

$$= \frac{4100}{8000} = \frac{41}{80} = 0.5125$$

EBrev
(131)

5) During a Year 1, 8000 policies had the above claim experience. Assume the number of claims followed a conditional Poisson(λ) distribution. Assume $\hat{\sigma}_V^2 = 0.5874$. A randomly selected policyholder had 1 claim in Year 1. Determine the semiparametric empirical Bayes estimate of the number of claims in Year 2 for the same policyholder.

$$\hat{a} = \hat{\sigma}_V^2 - \hat{V} = 0.5874 - 0.5125 = 0.0749$$

$$\hat{z} = \frac{1}{1 + \frac{\hat{V}}{\hat{a}}} = \frac{1}{1 + \frac{0.5125}{0.0749}} = 0.1275$$

$$P_C^1 = \bar{X} + \hat{z} (x_i - \bar{X}) =$$

$\rightarrow K = 6.8425$

$$0.5125 + 0.1275(1 - 0.5125) = \boxed{0.5747}$$

6) Suppose that the claim amount X is uniform($0, \Theta$) where the prior distribution of Θ has pdf $\pi(\theta) = \frac{500}{\theta^2}$ for $\theta > 500$. Two independent claims of 400 and 600 are observed. Find the pdf of the posterior distributions of Θ . Hint: $f(x|\theta) = f(400, 600|\theta) = \frac{1}{\theta\theta}$ for $\theta > \max(600, 400) = 600$. Hence the support of the posterior pdf is $\theta > 600$.

$$\pi(\theta|x) \propto \pi(\theta) f(x|\theta) \propto \theta^{-4}$$

$$\int_{600}^{\infty} C \theta^{-4} d\theta = C \frac{\theta^{-3}}{-3} \Big|_{600}^{\infty} = C \frac{1}{3(600)^3} = 1$$

$$\text{So } \pi(\theta|x) = 3(600)^3 \theta^{-4} = 648,000,000 \theta^{-4} \text{ for } \theta > 600.$$

$$(\text{=} 6.48 (10^8) \theta^{-4})$$

NOT ON 2024 exam 3

$$\bar{x} = \frac{9+4+8}{3} = 7$$

7) Suppose the number of claims per year for a policyholder follows a Poisson(λ) distribution where $\lambda \sim \text{gamma}(\alpha = 2.5, \theta = 4)$. There were 9, 4, and 8 claims in years 1 through 3 for this policyholder.

$$n\bar{x} = \sum x_i = 21, \quad \bar{x} = 7$$

(Note: you can get the posterior and predictive distributions from Table 51.1 on the back of p.9 of the exam 3 review.)

132a) was good too 142

→ a) Determine the Bayesian credibility estimate (posterior mean) for this policyholder's expected claim frequency in year 4. The answer is $\alpha^* \theta^*$ where $\theta^* = \frac{1}{\gamma^*}$.

posterior $\sim \text{Gamma}(\alpha^*, \theta^*)$ so

$$\theta^* = \frac{1}{\gamma^*}$$

$$\gamma^* = 3.25$$

$$\text{posterior mean} = \alpha^* \theta^* =$$

$$(\alpha + n\bar{x}) \left(\frac{1}{\frac{1}{\theta} + n} \right) = \left(2.5 + 3 \frac{21}{3} \right) \left(\frac{1}{\frac{1}{4} + 3} \right)$$

$$= 23.5 (0.3077) = 23.5 \left(\frac{1}{3.25} \right) = \boxed{7.2308}$$

b) Determine the Bayesian premium for year 4 (the mean of the predictive distribution).

$$\text{predictive} \sim \text{NB} (r = \alpha^*, \beta = \frac{1}{\gamma^*})$$

$$\text{Bayesian premium} = \alpha^* \frac{1}{\gamma^*} = \alpha^* \theta^* = \boxed{7.2308}$$

$$132a) \alpha' = \alpha + k = \alpha + n\bar{x} = \alpha^*, \quad \theta' = \frac{\theta}{1 + n\theta} = \theta^*$$

Table 51.2) $\alpha^* = \alpha + n\bar{x}$, $\gamma^* = \gamma + n$, $\gamma = \frac{1}{\theta}$, $\gamma^* = \frac{1}{\theta^*}$