

CR2) \downarrow A random variable is a real valued function for which the domain is a sample space.

ex) flip fair coin twice $S = \{HH, HT, TH, TT\}$.

Let $\mathbb{I} = \#$ of heads.

$$\mathbb{I}(HH) = 2, \mathbb{I}(HT) = \mathbb{I}(TH) = 1, \mathbb{I}(TT) = 0.$$

usually do not use this notation

	g	0	1	2
$P(\mathbb{I}=g)$		$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$

} use this notation

2) The event $\{\mathbb{I}=g\} = \{E_i \in S : \mathbb{I}(E_i) = g\}$.

Hence $P(\mathbb{I}=g)$ is the sum of all sample points in S that are assigned the value g .

ex) p23
$$\mathbb{I}_E = \begin{cases} 1 & \text{event } E \text{ occurs} \\ 0 & \text{event } E \text{ does not occur} \end{cases}$$

\mathbb{I}_E is the indicator RV for event E .

p24 3) know for final The cumulative distribution function

(cdf) of a RV X is $F(x) = P(X \leq x)$ for $x \in (-\infty, \infty) = \mathbb{R}$.

1) If $F(x)$ is a cdf, then

df 1) $F(-\infty) = \lim_{x \rightarrow -\infty} F(x) = 0$

df 2) $F(\infty) = \lim_{x \rightarrow \infty} F(x) = 1$

df 3) $F(x)$ is a nondecreasing function of x ! if

$x_1 < x_2$, then $F(x_1) \leq F(x_2)$.

(not in book) defn) F(x) is right continuous $x \downarrow x_0 \Rightarrow F(x) = F(x_0)$ for any x_0

13.5

Note: if $A = \{\bar{X} < x_1\}$ and $B = \{\bar{X} < x_2\}$ and $x_1 < x_2$

then $A \subseteq B$, Also $\{Y \leq y_1\} = \{\omega \in S: Y(\omega) \leq y_1\}$

ex 2.2-2.4

5) * A RV is discrete if it can only assume a finite or countably infinite number of distinct values. The collection of these probabilities is the probability distribution of the discrete RV.

6) * For a discrete RV Y , the probability mass function (pmf) $P(y) = P(Y=y)$ for all y . Generally if $P(y) = 0$, that value of y is omitted.

7) For the function $P(y)$ to be a pmf for a discrete RV, the following must hold

- i) $0 \leq P(y) \leq 1$ $\forall y$ for all
- ii) $\sum_y P(y) = 1$ where the sum is over all values $y: P(y) > 0$.

8) common E1 problem give a table with a $P(y)$ omitted.

ex)

y	0	1	2	3
$P(y)$.1	.2	.3	

Then $P(3) = .4$
 $= 1 - .1 - .2 - .3$

9) A probability distribution is a model or hopefully useful approximation for the data population.

that we want information about. (14)

ex) POP = IQ scores of college graduates
SIO " "

10) Know p34 Let Y be a discrete RV with pmf $P(y)$. Then the expected value of Y is

$$E(Y) = \sum_{\substack{y \\ y: P(y) > 0}} y \cdot P(y).$$

11) common problem Given a table of y and $P(y)$, find $E(Y)$.

ex)

y	0	1
$P(y)$	$\frac{1}{2}$	$\frac{1}{2}$

 $E(Y) = \sum y P(y) = 0(\frac{1}{2}) + 1(\frac{1}{2}) = \boxed{\frac{1}{2}}$

ex)

y	-1	1
$P(y)$	$\frac{1}{2}$	$\frac{1}{2}$

 $E(Y) = -1(\frac{1}{2}) + 1(\frac{1}{2}) = \boxed{0}$

ex)

y	-2	-1
$P(y)$	$\frac{1}{2}$	$\frac{1}{2}$

 $E(Y) = -2(\frac{1}{2}) + -1(\frac{1}{2}) = \boxed{-\frac{3}{2}}$

2) $E(Y) = \mu_{mv}$ = population mean is a measure of location of the POP

ex) mean ht of adult US women
wt men
network
yearly income

and if $g(Y)$ is a real valued function of Y , then $g(Y)$ is a RV and $E[g(Y)] = E[\bar{g}(Y)] = \sum_y g(y) P(y)$ (14.5)

14) know p40 The variance of a RV Y is

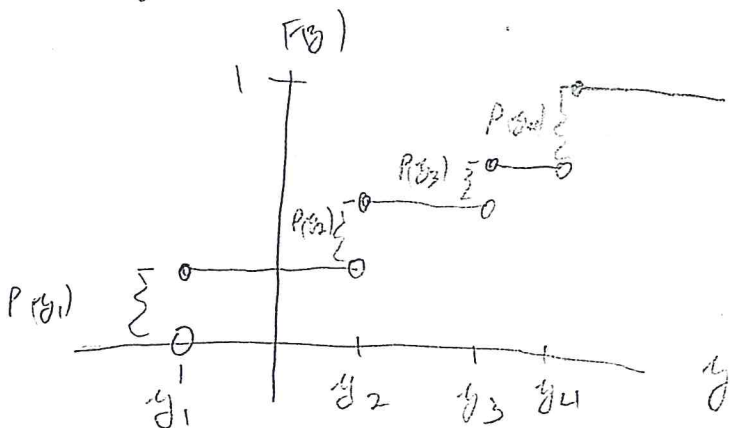
$V(Y) = \text{var}(Y) = E(Y - E(Y))^2$. The standard deviation (SD) of Y is $SD(Y) = \sqrt{V(Y)}$. know that $V(Y) \geq 0$ and $SD(Y) \geq 0$.

15) $V(Y) = \sigma^2$, $SD(Y) = \sigma$
sigma

$$V(Y) = \sum (y - E(Y))^2 P(y) \text{ so } g(y) = (y - E(Y))^2$$

16) Common final problem: Given a table of g and $P(y)$, find $E[\bar{g}(Y)]$, $V(Y)$ and $SD(Y)$.

17) If Y has pmf $P(y)$, then the cdf of Y is a step function with jumps of size $P(y)$ for values of y where $P(y) > 0$. See HW3 #3.



$$P(Y=y) = F(y) - F(y-1) = P(Y=y)$$

for $y: P(y) > 0$

y	y_1	y_2	y_3	y_4
$P(y)$	$P(y_1)$	$P(y_2)$	$P(y_3)$	$P(y_4)$

ex) $P(y) = 1$ $\frac{P(y-E(Y))}{1}$ is

$E(Y) = 7(1) = 7$

$V(Y) = 0, SD(Y) = \sqrt{0} = 0$

ex) Quiz 3 may ask you to make a table to compute σ or σ^2

y	P(y)	y P(y)	y - E(Y)	(y - E(Y)) ² P(y)
3	.3	.9	-1	0.3
4	.4	1.6	0	0
5	.3	1.5	1	0.3

$\sum P(y) = 1$ $E(Y) = 4$

$0.6 = V(Y)$ =

$\sum (y - E(Y))$ is not useful but

$\sum (y - E(Y))^2 P(y)$

$\sum (y - E(Y)) P(y) = E(Y - E(Y)) = 0$

$\sigma = SD(Y) = \sqrt{0.6} = 0.7746$

ex)

y	P(y)	y P(y)	y - E(Y)	(y - E(Y)) ² P(y)
1	.4	.4	-3	3.6
2	.1	.2	-2	0.4
6	.3	1.8	2	1.2
8	.2	1.6	4	3.2

$\sum P(y) = 1$ $E(Y) = 4$

$6.4 = V(Y)$ = $\sum (y - E(Y))^2 P(y)$

$\sigma = SD(Y) = \sqrt{6.4} = 2.52983$

After making a table or two, you may want to

say $\sigma^2 = V(Y) = \sum (y - E(Y))^2 P(y) =$

$(-4)^2(.4) + (-2)^2(.1) + (6-4)^2(.3) + (8-4)^2(.2) = 6.4$

Also short cut formula $V(Y) = E(Y^2) - (E(Y))^2$ is easier,

of the population measure of spread (variance)

15.5

19) Empirical rule! For many bell shaped populations,

$\mu \pm \sigma$ contains $\approx 68\%$ of measurements

$\mu \pm 2\sigma$ 95%

$\mu \pm 3\sigma$ 99.7% \approx all



20) * Expectation rules:

$E(c) = c$ if $g(Y) \equiv c$ is a constant,

$$E[\bar{c}g(Y)] = c E[\bar{g}(Y)]$$

$$E\left[\sum_{i=1}^k g_i(Y)\right] = \sum_{i=1}^k E[g_i(Y)]$$

expected value of a sum is the sum of the expected values

21) ^{PH1} know short cut formula for variance

$$V(Y) = E(Y - EY)^2 = E(Y^2) - [E(Y)]^2$$

proof) Let $\mu = EY$, $V(Y) = E(Y - \mu)^2 = E[Y^2 - 2Y\mu + \mu^2]$

$$= EY^2 - 2\mu E(Y) + \mu^2 = EY^2 - 2\mu^2 + \mu^2$$

$$= EY^2 - \mu^2$$

22) know $E(Y^2) = V(Y) + [E(Y)]^2$

y	$P(y)$	$yP(y)$	$y^2P(y)$
1	.4	.4	.4
2	.1	.2	.4
6	.3	1.8	10.8
8	.2	1.6	12.8
		$E(Y) = 4$	$24.4 = E(Y^2)$

$$V(Y) = E(Y^2) - [E(Y)]^2 = 24.4 - 16 = 8.4 \text{ as before.}$$

$$SD(Y) = \sigma = \sqrt{8.4}. \quad \text{Also}$$

$$E(Y^2) = \sum y^2 P(y) = 1^2(.4) + 2^2(.1) + 6^2(.3) + 8^2(.2) = 24.4$$

23} * P 26 An expt is a binomial experiment if

- 1) the expt consists of n identical trials
- 2) each trial has 2 outcomes = "success" and F
- 3) For each trial, $P(\text{Success}) = p$, $P(F) = 1-p$.
- 4) The trials are independent.
- 5) RV $Y =$ number of successes in n trials

Note: "Success" does not have the dictionary meaning, it is what you count, eg males, females
 Number of people that die within 5 years of being diagnosed with colon cancer.

ex) Flip coin n times, count # of heads

Y counts # for Clinton. This expt is approx binomial (the trials are approx ind if the sample size $n \ll N = \text{pop size}$).

Similarly randomly selected 1000 people, Y counts # that test HIV positive

ex) A voluntary response sample consists of people who choose themselves by responding to a general appeal. People with strong opinions are likely to respond and the sample does not represent a pop and not ind.

verbally

so not a binom expt. eg Ann Landers asked her readers if they had it to do over again would they have kids (after getting a letter from a parent with horrible kids). Almost 10000 parents responded and 70% said no.

american idol
the voice
non-scientific poll

In a random sample \approx binom expt, 9% would say no.

ex) prof needs student to comment on his teaching and picks several A students, $Y = \#$ that give high rating. This is not a bin expt (pop of prof students is not represented, not ind)

4) know RV Y has a binomial distribution based on n trials with success prob p , written $Y \sim \text{bin}(n, p)$ if the pmf

$\binom{n}{y} p^y (1-p)^{n-y}$ for $y = 0, 1, \dots, n$, The pmf formula is for $p \in (0, 1)$.
 Note: $q = 1-p$. If there are n trials and y s's then there are $n-y$ F's.

$$P(y \text{ s's and } n-y \text{ F's}) = p^y (1-p)^{n-y} \text{ by ind.}$$

There are $\binom{n}{y}$ sequences with y s's. By

$$\text{the binomial th, } 1 = (p+1-p)^n = \sum_{y=0}^n \binom{n}{y} p^y (1-p)^{n-y}$$

So $P(y)$ is a pmf.

25} know! If $Y \sim \text{b.i.n}(n, p)$, $E(Y) = np$ and $V(Y) = np(1-p)$.

common error! try to find $E(Y) = \sum y P(y)$ and

$V(Y) = \sum (y - EY)^2 P(y)$ when $Y \sim \text{b.i.n}(n, p)$. The above formulas are much easier.

26} know A RV X is a Bernoulli RV if $X \sim \text{b.i.n}(n=1, p)$

Note: p 35 proof that $E(Y) = np$ for $Y \sim \text{b.i.n}(n, p)$

$$E(Y) = \sum y P(y) = \sum_{y=0}^n y \frac{n!}{y!(n-y)!} p^y (1-p)^{n-y} =$$

$$np \sum_{y=1}^n \frac{(n-1)!}{(y-1)!(n-y)!} p^{y-1} (1-p)^{n-y}$$

$y=0$ term is 0
 $y=1$
 $y! = 0!$
 $n! = n \cdot (n-1)!$

Let $z = y-1$ then $y = n \rightarrow z = n-1$
 $y = 1 \rightarrow z = 0$ } like u substitution

trick

$$\text{and } E(Y) = np \sum_{z=0}^{n-1} \frac{(n-1)!}{z!(n-1-z)!} p^z (1-p)^{n-1-z} =$$

iii' $\sum_{z=0}^{\infty} 1 = \sum_{z=0}^{\infty} P(Z=z)$ \leftarrow $\sum_{z=0}^{\infty} P(Z=z)$ iii'

175

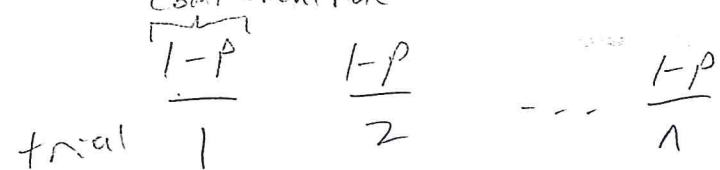
ex) Suppose $Y \sim \text{bin}(n, p)$ where Y counts the number of times where outcome D occurred where $P(D) = p$ for a single experiment.

be for
for ex
and 1's

- Find
- i) $P(D \text{ occurred in none of the } n \text{ expts})$
 - ii) $P(D \text{ occurred in at least one of the } n \text{ expts})$
 - iii) $P(D \text{ occurred in all } n \text{ expts})$
 - iv) $P(D \text{ occurred in not all } n \text{ expts})$



i) $P(\text{none}) = P(Y=0) = \binom{n}{0} p^0 (1-p)^n = (1-p)^n$



$Y \in \{0, 1, \dots, n\}$

ii) $P(\text{at least one}) = P(Y \geq 1) = 1 - P(Y < 1) = 1 - P(Y=0)$

$= P(\text{not none}) = 1 - P(\text{none}) = 1 - (1-p)^n = 1 - \binom{n}{0} p^0 (1-p)^n$

iii) $P(\text{all}) = P(Y=n) = \binom{n}{n} p^n (1-p)^0 = p^n$



iv) $P(\text{not all}) = 1 - P(\text{all}) = 1 - P(Y=n) = 1 - \binom{n}{n} p^n (1-p)^0 = 1 - p^n$

27) $P(Y \text{ is at most } k) = P(Y \leq k), \quad P(Y \text{ is at least } k) = P(Y \geq k)$

$$P(Y \text{ is at most } k) = P(Y \leq k) = P(0) + \dots + P(k)$$

$$= 1 - P(Y > k) = 1 - P(Y \geq k+1) = 1 - [P(k+1) + P(k+2) + \dots + P(n)]$$

$$P(Y \text{ is at least } k) = P(Y \geq k) = P(k) + \dots + P(n)$$

$$= 1 - P(Y < k) = 1 - P(Y \leq k-1) = 1 - P(0) - \dots - P(k-1)$$

ex) common E1 problem: Fair die is rolled 3 times.

Y counts # of 5's. Then $Y \sim \text{b.n.}(n=3, p=\frac{1}{6})$

Y	0	1	2	3
$P(Y)$.5787	.3472	.0694	.0046

sums to 0.9999 due to rounding

a) $E(Y) = np = 3 \cdot \frac{1}{6} = \boxed{0.5}$

b) $V(Y) = np(1-p) = 3 \cdot \frac{1}{6} \cdot \frac{5}{6} = \frac{5}{12} = \boxed{0.4167}$

sum to 1

c) $P(\text{none of the rolls are 5's}) = P(Y=0) = P(0) = \binom{3}{0} \left(\frac{1}{6}\right)^0 \left(\frac{5}{6}\right)^3$
 $= \left(\frac{5}{6}\right)^3$

sum to 1

d) $P(\text{at least one roll is a 5}) = P(Y \geq 1) = P(1) + P(2) + P(3)$
 $= 1 - P(0) = 1 - \left(\frac{5}{6}\right)^3$

e) $P(\text{all 3 rolls are 5's}) = P(Y=3) = P(3) = \binom{3}{3} \left(\frac{1}{6}\right)^3 \left(\frac{5}{6}\right)^0$
 $= \left(\frac{1}{6}\right)^3$

f) $P(\text{not all 3 rolls are 5's}) = P(Y \neq 3) = 1 - P(Y=3)$
 $= 1 - P(3) = P(0) + P(1) + P(2) = 1 - \left(\frac{1}{6}\right)^3$

$$= \binom{3}{0} \left(\frac{1}{6}\right)^0 \left(\frac{5}{6}\right)^3 + \binom{3}{1} \left(\frac{1}{6}\right)^1 \left(\frac{5}{6}\right)^2 = \left(\frac{5}{6}\right)^3 + 3 \frac{1}{6} \left(\frac{5}{6}\right)^2 =$$

$$0.5787 + 0.3472 = \boxed{0.9259}$$

$$h) P(\text{at least 2 rolls are 5's}) = P(Y \geq 2)$$

$$= P(2) + P(3) = 1 - P(0) - P(1) = 1 - .9259 = \boxed{0.0741}$$

28) * Common error: $q = 1 - p$ is given in the story problem and student uses $Y \sim \text{bin}(n, q)$ instead of $Y \sim \text{bin}(n, p)$, see HW 3 #6.

29) * P29-30 A poisson(λ) RV Y is used to model the total # of occurrences of some phenomenon in a fixed period of time or a fixed region of space where λ is the average value of Y .

eg # telephone calls in a 5 minute interval (at a center with several phones)

radioactive particles that hit a target in 3 hours

30) P29 know A RV Y has a poisson dist

$$\text{if } Y \text{ has pmf } P(y) = \frac{\lambda^y e^{-\lambda}}{y!}, \quad y = 0, 1, 2, \dots$$

for $\lambda > 0$.

31) know P36 If $Y \sim \text{Poisson}(\lambda)$ then $E(Y) = V(Y) = \lambda$,
(distributions)