

1) Let  $Y_1, \dots, Y_n$  be iid Poisson( $\theta$ ) random variables.

20 a) Find the method of moments estimator for  $\theta$ .

$E(Y) = \theta \stackrel{\text{set}}{=} \bar{Y}$  so  $\hat{\theta} = \bar{Y}$

20 b) Find the method of moments estimator for  $1/\theta$ .

$g(\theta) = \frac{1}{\theta}$  so  $\widehat{\left(\frac{1}{\theta}\right)} = \frac{1}{\bar{Y}} = g(\bar{Y})$   
by MM Theorem

5 → c) The Fisher information  $I_1(\theta) = 1/\theta$ . Find  $I_n(1/\theta) = \frac{n I_1(\theta)}{[\tau'(\theta)]^2} = \frac{n}{\theta \theta^4}$

$\tau(\theta) = \theta^{-1}, \tau'(\theta) = -\theta^{-2}$   
 $= \boxed{n\theta^3}$

10 d) Find the FCRLB for unbiased estimators of  $\tau(\theta) = 1/\theta$ .

$\frac{\theta^{-4}}{n \cdot \frac{1}{\theta}} = \frac{[\tau'(\theta)]^2}{n I_1(\theta)} = \boxed{\frac{1}{n\theta^3}} = \frac{1}{I_n(\theta)}$

5 → e) The UMVUE for  $e^{a\theta}$  is  $U_n = \exp[\sum_{i=1}^n Y_i \log(\frac{a+Y_i}{n})]$ . Do you expect the UMVUE  $U_n$  to achieve the FCRLB for unbiased estimators of  $e^{a\theta}$ ? Explain briefly.

NO,  $T_n = \sum_{i=1}^n Y_i$  is the complete sufficient statistic and  $U_n$  is a nonlinear function of  $T_n$ .

Answered question for (e)

60  $U_n = \left(\frac{a+T_n}{n}\right)^{T_n}$  is better than don't need  $a > -1$

2) Let  $Y_1, \dots, Y_n$  be a random sample from a distribution with pdf

$$f(y) = \frac{2y}{\theta^2}, \quad 0 < y < \theta.$$

Then  $E(Y_i) = 2\theta/3$  and  $V(Y_i) = \theta^2/18$ . Let  $T = c\bar{Y}$  be an estimator of  $\theta$  where  $c$  is a constant. Find the mean square error (MSE) of  $T$  as a function of  $c$  (and of  $\theta$  and  $n$ ).

$$\begin{aligned} E(T) &= cE(\bar{Y}) = c \frac{2}{3}\theta, \quad V(T) = c^2 V(\bar{Y}) = c^2 \frac{V(Y)}{n} = \frac{c^2 \theta^2}{18n} \\ \text{MSE}(T) &= V(T) + [\text{bias}(T)]^2 = \frac{c^2 \theta^2}{18n} + \left(\frac{c}{3}\theta - \theta\right)^2 \\ &= \frac{c^2 \theta^2}{18n} + \frac{\theta^2}{9} (2c-3)^2 = \frac{c^2 \theta^2}{18n} + \theta^2 \left(\frac{2c}{3} - 1\right)^2 \\ &= \frac{c^2 \theta^2}{18n} + \frac{\theta^2}{9} (2c-3)^2 \end{aligned}$$

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3) Let  $X_1, \dots, X_n$  be independent identically distributed random variables with pdf

$$f(x) = \frac{1}{\lambda} \phi x^{\phi-1} \frac{1}{1+x^\phi} \exp\left[-\frac{1}{\lambda} \log(1+x^\phi)\right]$$

where  $x, \phi$ , and  $\lambda$  are all positive. ASSUME  $\phi$  IS KNOWN. Find the uniformly minimum unbiased estimator (UMVUE) of  $\lambda$ .

Hint: You may use the fact that  $\log(1+X_i^\phi) \sim \text{Gamma}(\nu=1, \lambda)$ .

1) PRET  $T = \sum_{i=1}^n \log(1+X_i^\phi) \sim G(n, \lambda)$  is complete

$$ET = n\lambda = \sum_{i=1}^n E[\log(1+X_i^\phi)] = \sum_{i=1}^n \lambda = n\lambda$$

so  $\frac{T}{n} = \frac{\sum_{i=1}^n \log(1+X_i^\phi)}{n}$  is the UMVUE by LSU