

$\sum X_i \sim \text{bin}(n, p)$

$$d \left[p - \frac{1}{n^2} (n p(1-p) + n^2 p^2) \right] = d \left[p - \frac{p}{n} + \frac{p^2}{n} - p^2 \right]$$

$$d \left[\frac{n-1}{n} p + p^2 \frac{1-n}{n} \right] = d \frac{n-1}{n} (p - p^2) = d \frac{n-1}{n} p(1-p)$$

So take $d = \frac{n}{n-1}$ and $U = \frac{n}{n-1} \bar{X} (1-\bar{X})$ is the UMVUE.

method iii)

$W(X) = X_1 (1-X_2)$ has $E(W(X)) = EX_1 - EX_1 EX_2 = p - p^2 = p(1-p)$. So

$Eg [X_1 (1-X_2) | \sum X_i = u] = p [X_1 (1-X_2) = 1 | \sum_{i=1}^n X_i = u]$

(trick: $X_1 (1-X_2) \in \{0, 1\}$, $X_1 (1-X_2) = 1$ iff $X_1 = 1, X_2 = 0$ so Bernoulli)

$E(Y) = P(Y=1)$ if Y is Bern

$$\frac{P_p(X_1=1, X_2=0, \sum_{i=3}^n X_i = u-1)}{P_p(\sum_{i=1}^n X_i = u)} = \frac{\binom{n-2}{u-1} p^{u-1} (1-p)^{n-2-u+1}}{\binom{n}{u} p^u (1-p)^{n-u}} = \frac{\binom{n-2}{u-1}}{\binom{n}{u}} = \frac{(n-2)!}{(u-1)! (n-1-u)!} \cdot \frac{u! (n-u)!}{n!}$$

$$= \frac{(n-2)! \cdot u(u-1)! (n-u)(n-1-u)!}{n(n-1)(n-2)! (u-1)! (n-1-u)!} = \frac{u(n-u)}{n(n-1)}$$

So the UMVUE is $T(X) = \frac{\sum X_i (n - \sum X_i)}{n(n-1)} = \frac{n}{n-1} \bar{X} (1-\bar{X})$.

ch 7 1] p183 A hypothesis is a statement about a population parameter.

2] p. 183 In hypothesis testing there are 2 competing hypotheses: the null hypothesis H_0 and the alternative hypothesis $H_1 \equiv H_A$.

3) Let $\theta \in \Theta$. $H_0: \theta \in \Theta_0$ $H_1: \theta \in \Theta_1$

$\Theta_0 \cap \Theta_1 = \emptyset$, $\Theta_0, \Theta_1 \subseteq \Theta$.

4) p. 183 A hypothesis test is a rule for rejecting H_0 .

ex] p 184 Often there is a rejection region R .
Reject H_0 if test statistic $T(\underline{y}) \in R$, otherwise fail to reject H_0 .

ex] $Y \sim N(\mu, \frac{1}{9})$, $\mu \in \{0, 1\}$, $\sigma^2 = \frac{1}{9}$, $3\sigma = 1$
 $H_0: \mu = 0$ $H_1: \mu = 1$ Reject H_0 if $Y \geq \frac{1}{2}$
otherwise fail to reject H_0 .



		Decision
	reject H_0	fail to reject H_0
H_0	type I error	
H_1		type II error

A type I error is rejecting H_0 when H_0 is true
A type II error is failing to reject H_0 when H_0 is false.

6] know P. 183 The power function of a hypothesis test is $\beta(\theta) = P_{\theta} (H_0 \text{ is rejected})$ for $\theta \in \Theta$.

Want $\beta(\theta) \approx 0$ for $\theta \in \Theta_0$ and $\beta(\theta) \approx 1$ for $\theta \in \Theta_1$.

ex] $\beta(\theta) = P_{\theta} (T(\underline{Y}) \in R)$ if there is a test statistic $T(\underline{Y})$ with rejection region R .

Fact] For $\theta \in \Theta_0$, $\beta(\theta) = P_{\theta}(\text{type I error})$. For $\theta \in \Theta_1$, $\beta(\theta) = 1 - P_{\theta}(\text{type II error})$.

ex] $Y \sim N(\mu, \frac{1}{9})$ $\mu \in \{0, 1\} = \Theta$, $\Theta_0 = \{0\}$, $\Theta_1 = \{1\}$ 530 45

$$B(0) = P_0(Y > \frac{1}{2}) \stackrel{Y \sim N(0, \frac{1}{9})}{=} P\left(\frac{Y-0}{\frac{1}{3}} > \frac{\frac{1}{2}-0}{\frac{1}{3}}\right) = P(Z > 1.5)$$

$$= 1 - P(Z < 1.5) = 1 - .9332 = .0668$$

\uparrow \uparrow
 $N(0,1)$ table

$$B(1) = P_1(Y > \frac{1}{2}) \stackrel{Y \sim N(1, \frac{1}{9})}{=} P\left(\frac{Y-1}{\frac{1}{3}} > \frac{\frac{1}{2}-1}{\frac{1}{3}}\right) = P(Z > -1.5) = .9332$$

7] know p184 For $0 \leq \alpha \leq 1$, a test with power function $B(\theta)$ is a size α test if $\sup_{\theta \in \Theta_0} B(\theta) = \alpha$ and a level α test if $\sup_{\theta \in \Theta_0} B(\theta) \leq \alpha$.

Tip: Often $\Theta_0 = (a, \theta_0]$, $\Theta_1 = (\theta_0, b)$ and $\alpha = \beta(\theta_0) = P_{\theta_0}$ (type I error).

ex] The last ex is a level 0.07 test and a size .0668 test.

§7.2 8) p185 Consider all level α tests of $H_0: \theta \in \Theta_0$ vs $H_1: \theta \in \Theta_1$. A uniformly most powerful (UMP) level α test is a level α test with power function $B_{ump}(\theta) \geq \beta(\theta) \forall \theta \in \Theta_1$ where β is the power function for any level α test of H_0 vs H_1 .

9) know Neyman Pearson NP Lemma: Consider testing $H_0: \theta = \theta_0$ vs $H_1: \theta = \theta_1$ where the pdf or pmf

45.5
corresponding to θ_i is $f(\underline{y}|\theta_i)$ for $i=0,1$.

The level α UMP test

rejects H_0 if $f(\underline{y}|\theta_1) > k f(\underline{y}|\theta_0)$

and rejects H_0 with prob γ if $f(\underline{y}|\theta_1) = k f(\underline{y}|\theta_0)$
for some $k \geq 0$.

$$\alpha = \beta(\theta_0) = P_{\theta_0} [f(\underline{y}|\theta_1) > k f(\underline{y}|\theta_0)] + \gamma P_{\theta_0} [f(\underline{y}|\theta_1) = k f(\underline{y}|\theta_0)].$$

Note) With p.d.f.s, $\gamma = 0$ and the rejection region $\underline{y} \in R$ is $f(\underline{y}|\theta_1) > k f(\underline{y}|\theta_0)$.

10] Know There is a k where you reject H_0
if $\frac{f(\underline{y}|\theta_1)}{f(\underline{y}|\theta_0)} > k$. Put the test into

"useful form" by simplifying so
that you reject H_0 if $T(\underline{y}) > d$ (if the
ratio is an increasing function of $T(\underline{y})$) or
if $T(\underline{y}) < c$ (if ratio is a decreasing function of $T(\underline{y})$)

11] p186 Assume $\sup_{\theta \in \Theta_0} \beta(\theta) = \beta(\theta_0)$

Know Test $H_0 \theta \leq \theta_0$ vs $H_1 \theta > \theta_0$ } right
or $H_0 \theta < \theta_0$ vs $H_1 \theta \geq \theta_0$ } tail

or $H_0 \theta \geq \theta_0$ vs $H_1 \theta < \theta_0$ } left
or $H_0 \theta > \theta_0$ vs $H_1 \theta \leq \theta_0$ } tail.

Pick $\theta_1 \in \Theta_1$ and use NP to find the
UMP level α test for $H_0^* \theta = \theta_0$ vs $H_1^* \theta = \theta_1$.
So reject H_0^* if $f(\underline{y}|\theta_1) > k f(\underline{y}|\theta_0)$ and
reject H_0^* with prob γ if $f(\underline{y}|\theta_1) = k f(\underline{y}|\theta_0)$
for some $k \geq 0$ where $\alpha = \beta(\theta_0)$ in q . This test

is the UMP level α test for $H_0 \theta \in \Theta_0$ vs $H_1 \theta \in \Theta_1$,
if k does not depend on the value of $\theta_1 \in \Theta_1$.
(So k depends only on θ_0 and α .)

12] Know p.186 one sided tests for exp families.

Let Y_1, \dots, Y_n be iid with pdf or pmf
 $f(y|\theta) = h(y)c(\theta)\exp[w(\theta)t(y)]$

where $w(\theta)$ is increasing (eg $w'(\theta) > 0$).

Let $T(\underline{y}) = \sum_{i=1}^n t(y_i)$. I) Let $\theta_1 > \theta_0$

Consider $H_0 \theta = \theta_0$ vs $H_1 \theta = \theta_1$
 $H_0 \theta = \theta_0$ vs $H_1 \theta > \theta_0$
 $H_0 \theta \leq \theta_0$ vs $H_1 \theta > \theta_0$.

The UMP level α test rejects H_0 if
 $T(\underline{y}) > k$ and rejects H_0 with prob γ if $T(\underline{y}) = k$
where $\alpha = P_{\theta_0}(T(\underline{Y}) > k) + \gamma P_{\theta_0}(T(\underline{Y}) = k)$.

I >
goes
with
 H_1
 $\theta_1 > \theta_0$

II

Let $\theta_1 < \theta_0$.

consider $H_0 \theta = \theta_0$ vs $H_1 \theta = \theta_1$
 $H_0 \theta = \theta_0$ vs $H_1 \theta < \theta_0$
 $H_0 \theta \geq \theta_0$ vs $H_1 \theta < \theta_0$.

The UMP level α test rejects H_0 if
 $T(\underline{y}) < k$ and rejects H_0 w.p γ if $T(\underline{y}) = k$
where $\alpha = P_{\theta_0}(T(\underline{Y}) < k) + \gamma P_{\theta_0}(T(\underline{Y}) = k)$.

II <
goes
with
 H_1
 $\theta_1 < \theta_0$

13]

Mnemonic p.199 The ">" or "<" goes
with H_1 and " $\theta_1 > \theta_0$ " or " $\theta_1 < \theta_0$ "

The inequality in the rejection region is the same as
the inequality in the alternative hypothesis.

14) know suppose you find

$$f(y|\theta) = h(y) c(\theta) \exp [\tilde{w}(\theta) \tilde{x}(y)]$$

where \tilde{w} is decreasing (eg $\tilde{w}'(\theta) < 0$).

Then set $w(\theta) = -\tilde{w}(\theta)$ and $x(y) = -\tilde{x}(y)$.

ex] See ex 7.2 - 7.5. See problems 7.10, 7.13, 7.14, 7.15, and 7.16 solns in Ch 12.

15) 6th problem for E3 final qual is finding UMP tests via NP lemma or exp fam theory.

ex] X_1, \dots, X_n iid $N(\mu, \sigma^2)$, $\sigma^2 > 0$ known. Find UMP level α test for $H_0: \mu = \mu_0$ vs $H_1: \mu > \mu_0$.

Soln method I) $f(x) = \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2}(\frac{x-\mu}{\sigma})^2} = \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2\sigma^2}(x^2 - 2\mu x + \mu^2)}$

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{x^2}{2\sigma^2}} e^{-\mu^2/2\sigma^2} \exp\left[\frac{\mu}{\sigma^2} x\right]$$

$w(\mu) \quad x(x)$

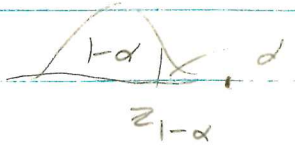
$w'(\mu) = \frac{1}{\sigma^2} > 0$ so $w \uparrow$. Let $P(Z \leq z_\alpha) = \alpha$ if $Z \sim N(0,1)$

so reject H_0 if $\sum_{i=1}^n X_i > k$ where

$$\alpha = P_{\mu_0}(\sum X_i > k) = P\left(\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} > z_{1-\alpha}\right)$$

$\bar{X} \stackrel{H_0}{\sim} N(\mu_0, \frac{\sigma^2}{n})$

test rejects H_0 if $\frac{\bar{X} - \mu_0}{\sigma/\sqrt{n}} > z_{1-\alpha}$



Method II) Use NP lemma: Fix $\mu_1 > \mu_0$.

$$\frac{f(x|\mu_1)}{f(x|\mu_0)} = \frac{e^{-\frac{1}{2\sigma^2} \sum (x_i - \mu_1)^2}}{e^{-\frac{1}{2\sigma^2} \sum (x_i - \mu_0)^2}} > K \quad \text{iff}$$

$$\frac{e^{-\frac{1}{2\sigma^2} \sum x_i^2} e^{-\frac{n\mu_1^2}{2\sigma^2}} e^{\frac{\mu_1}{\sigma^2} \sum x_i}}{e^{-\frac{1}{2\sigma^2} \sum x_i^2} e^{-\frac{n\mu_0^2}{2\sigma^2}} e^{\frac{\mu_0}{\sigma^2} \sum x_i}} > K$$

positive constant

iff $\exp\left(\frac{1}{\sigma^2} (\mu_1 - \mu_0) \sum x_i\right) > K'$

> 0 if $\mu_1 > \mu_0$ ← key step

iff $\sum x_i > K''$ where $P_{\mu_0}(\sum x_i > K'') = \alpha$

key step → Now K'' only depends on μ_0 and α NOT on $\mu_1 > \mu_0$. So the UMP test rejects H_0 if $\frac{\bar{x} - \mu_0}{\sigma/\sqrt{n}} > z_{1-\alpha}$.

ex) X_1, \dots, X_n iid $\text{Ber}(p)$, $p \in (0, 1)$, $f(p) = p^x (1-p)^{1-x}$
 Find UMP level α test for $H_0: p = p_0$ vs $H_A: p > p_0$.

Soln } $\left(f(x) = (1-p) \exp\left[\underbrace{\log\left(\frac{p}{1-p}\right)}_{\text{w.p.t}} x\right] \right)$ so reject H_0 if $\sum x_i > K$ and w.p. \forall if $\sum x_i = K$

omit } Fix $p_1 > p_0$. Now $\frac{f(x|p_1)}{f(x|p_0)} = \frac{p_1^{\sum x_i} (1-p_1)^{n-\sum x_i}}{p_0^{\sum x_i} (1-p_0)^{n-\sum x_i}}$

$$= \left(\frac{p_1(1-p_0)}{p_0(1-p_1)}\right)^{\sum x_i} \left(\frac{1-p_1}{1-p_0}\right)^n > K \quad \text{iff } \sum x_i \log\left(\frac{p_1(1-p_0)}{p_0(1-p_1)}\right) > K'$$

iff $\sum x_i > K''$, so reject H_0 if $\sum x_i > K$

$\frac{p_1 > p_0}{1-p_0 > 1-p_1}$

If $\rho_1 > \rho_0$ this test depends only on α and ρ_0 .
 So the UMP level α test rejects H_0 if $\sum_{i=1}^n X_i > c$
 and rejects H_0 w.p. γ if $\sum_{i=1}^n X_i = c$

(*) where $\alpha = P_{\rho_0} \left(\sum_{i=1}^n X_i > c \right) + \gamma P_{\rho_0} \left(\sum_{i=1}^n X_i = c \right)$.

Suppose $\rho_0 = \frac{1}{2}$ and $n=8$. Then $\sum_{i=1}^n X_i \stackrel{H_0}{\sim} \text{bin}(8, \frac{1}{2})$

c	8	7	6
$P_{\frac{1}{2}} \left(\sum_{i=1}^8 X_i = c \right)$	$\binom{8}{8} \left(\frac{1}{2}\right)^8 = \left(\frac{1}{2}\right)^8$	$\binom{8}{7} \left(\frac{1}{2}\right)^8 = 8 \left(\frac{1}{2}\right)^8$	$\binom{8}{6} \left(\frac{1}{2}\right)^8 =$
	$P_8 = 0.0039063$	$P_7 = 0.03125$	$P_6 = 0.10937$

If $\alpha = 0.05$, reject H_0 if $\sum X_i > 6$ and
 reject H_0 w.p. γ if $\sum X_i = \boxed{6=c}$ where by (*)

$$\alpha = P_8 + P_7 + \gamma P_6 \quad \text{or} \quad \gamma = \frac{\alpha - P_8 - P_7}{P_6}$$

$$= \frac{.05 - .0039063 - .03125}{.10937} = \frac{.0148437}{.10937} = .13572 = \gamma$$

p188 In general, (*) gives $\gamma = \frac{\alpha - P_{\rho_0}(\sum X_i > c)}{P_{\rho_0}(\sum X_i = c)}$

See HW11 #2

$$\int_1^{\infty} \theta x^{-\theta-1} dx = \theta \frac{x^{-\theta-1+1}}{-\theta-1+1} = -x^{-\theta} \Big|_1^{\infty} = -0 - (-1) = 1$$

ex) X_1, \dots, X_n iid $f(x|\theta) = \frac{\theta}{x^{\theta+1}}, x > 1$.

a) Find UMP level $\alpha = \text{size } \alpha$ test of $H_0: \theta = 1$ vs $H_1: \theta > 1$

soln I) $f(x|\theta) = I(x > 1) \theta \exp[-(\theta+1) \log x]$

$$= I(x > 1) \theta \exp\left[\underbrace{(\theta+1)}_{w(\theta)} \underbrace{[-\log(x)]}_{t(x)} \right]$$

$$T(x) = -\sum_{i=1}^n \log(X_i), \quad \text{UMP test rejects } H_0$$