

then $\prod_{i=1}^n I(y_i \in y^*)$ could be part of $g(\underline{T}(\underline{y})|\theta)$, 580 26

Typically y^* is an interval with endpoints a and b , not necessarily finite. For pdfs ^{and}

$$\prod_{i=1}^n I[y_i \in (a, b)] = \underbrace{I[a < y_{(1)} \leq y_{(n)} < b]}_{\text{i) use if } a \text{ and } b \text{ are unknown}} = \underbrace{I[a < y_{(1)}]}_{\text{ii) use if one of } a \text{ or } b \text{ is unknown}} I[y_{(n)} < b]$$

iii) use either i) or ii) if both a and b are known.

$$\prod_{i=1}^n I(y_i \in [a, b]) = I(a \leq y_{(1)} \leq y_{(n)} \leq b) = I[a \leq y_{(1)}] I[y_{(n)} \leq b]$$

$$\prod_{i=1}^n I[y_i \in (-\infty, b)] = I(y_{(n)} < b)$$

$$\prod_{i=1}^n I(y_i \in [a, \infty)) = I[y_{(1)} \geq a] \text{ etc.}$$

ex] Y_1, \dots, Y_n iid $U(\theta_1, \theta_2)$. Find a suff stat
ex 4.10 \underline{T} for (θ_1, θ_2) .

soln $f(\underline{y}) = \prod_{i=1}^n f(y_i) = \prod_{i=1}^n \frac{1}{\theta_2 - \theta_1} I[\theta_1 \leq y_i \leq \theta_2]$

$$= \frac{1}{(\theta_2 - \theta_1)^n} I[\theta_1 \leq y_{(1)} \leq y_{(n)} \leq \theta_2] \quad \left(\underbrace{1}_{h(\underline{y})} \right)$$

So $\underline{T}(\underline{y}) = (Y_{(1)}, Y_{(n)})$ by Factorization

If θ_1 is known, $f(\underline{y}) = \frac{1}{(\theta_2 - \theta_1)^n} I(y_{(n)} \leq \theta_2) \underbrace{I[\theta_1 \leq y_{(1)}]}_{h(\underline{y})}$
 $g(\underline{T}(\underline{y})|\theta_2)$

and $Y_{(n)}$ is suff for θ_2 .

If θ_2 is known, $f(\underline{y}) = \frac{1}{(\theta_2 - \theta_1)^n} I(\theta_1 \leq y_{(1)}) \underbrace{I(y_{(n)} \leq \theta_2)}_{h(\underline{y})}$
 $g(\underline{T}(\underline{y})|\theta_1)$

and $Y_{(1)}$ is suff for θ_1 . his Factorization

16] P114 If $\underline{T}(\underline{Y})$ is a sufficient statistic for $\underline{\theta}$, then any statistic $\underline{S}(\underline{Y})$ that is a one to one function of $\underline{T}(\underline{Y})$ (so $\underline{S}(\underline{Y}) = k(\underline{T}(\underline{Y}))$ where k is one to one) is also a sufficient statistic for $\underline{\theta}$.

ex] know Y_1, \dots, Y_n iid $N(\mu, \sigma^2)$ $\underline{T}(\underline{Y}) = (\sum Y_i, \sum Y_i^2)$ is sufficient for (μ, σ^2) , so $\underline{S}(\underline{Y}) = (\bar{Y}, S^2) = (\frac{1}{n} \sum Y_i, \frac{1}{n-1} (\sum Y_i^2 - n(\frac{\sum Y_i}{n})^2))$ is also suff for (μ, σ^2) .

17] P113-4 If $\underline{m} = \underline{T}(\underline{\theta})$ is one to one, then the parameter $\underline{\theta}$ has just been renamed and the distribution of Y_1, \dots, Y_n is not changed. Hence if $\underline{T}(\underline{Y})$ is sufficient for $\underline{\theta}$, then $\underline{T}(\underline{Y})$ is sufficient for \underline{m} .

proof of Th 4.5e

ex] one to one is crucial; If Y_1, \dots, Y_n are iid $N(\mu, \sigma^2)$, let $k(\sum Y_i, \sum Y_i^2) \equiv 1 = T'(\underline{Y})$. Then $T'(\underline{Y})$ is not sufficient for (μ, σ^2) since $\underline{Y} | T'(\underline{Y}) \sim \underline{Y} \sim N_n(\underline{\mu}, \sigma^2 I_n)$

which depends on $\underline{\theta} = (\mu, \sigma^2)$.

18] Know for final p111 A sufficient

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statistic $\underline{T}(Y)$ for $\underline{\theta}$ is a minimal sufficient statistic for $\underline{\theta}$ if

$\underline{T}(Y)$ is a function of $\underline{S}(Y)$ for any other suff stat $\underline{S}(Y)$. So
 $\underline{T}(Y) = K_S(\underline{S}(Y))$ for some function K_S .

19] Mnemonic: ^{memory aid} if Y_1, \dots, Y_n have pdf or pmf
 $f(y_1, \dots, y_n | \theta) = \underbrace{1}_{h(\underline{z})} \cdot \underbrace{\theta(T(\underline{z}) | \theta)}$ then

$\underline{S}(Y) = \underline{Y}$ is sufficient by factorization,
and $\underline{T} \equiv \underline{T}(Y)$ is a function of \underline{S} .

20] Want to use a minimal sufficient statistic which reduces the data as far as possible without losing information about $\underline{\theta}$.

Note] $\underline{T}_1(Y) = (Y_1, \dots, Y_n)$ and $\underline{T}_2(Y) = (Y_{(1)}, \dots, Y_{(n)})$ are sufficient, but \underline{T}_1 is not a function of \underline{T}_2 for $n \geq 1$. So \underline{T}_1 is not minimal suff.

21] Know p114 Any one to one function of a minimal sufficient statistic is a minimal sufficient statistic.

22) know for final p. 112 Suppose statistic 27.5

$\underline{T}(Y)$ has pdf or pmf $f(\underline{t}|\theta)$. Then

$\underline{T}(Y)$ is a complete sufficient statistic for θ iff

$$E_{\theta} [g(\underline{T}(Y))] = 0 \quad \forall \theta \in \Theta$$

implies that $P_{\theta} [g(\underline{T}(Y)) = 0] = 1 \quad \forall \theta \in \Theta$.

23) know p. 114 Bahadur's theorem Suppose

k does not depend on the sample size n .

Then a k dimensional complete suff stat

$\underline{T}(Y)$ is also minimal suff.

24) know for final Lehmann Scheffé LSM theorem!

Let iid sample \underline{y} have pdf or pmf $f(\underline{y}|\theta)$.

Let $c_{\underline{x}, \underline{y}}$ be a constant. If

there exists a function $\underline{T}(\underline{y})$ such

that for any two sample points \underline{x} and \underline{y} ,

the ratio $R(\theta) = \frac{f(\underline{x}|\theta)}{f(\underline{y}|\theta)} = c_{\underline{x}, \underline{y}} \quad \forall \theta \in \Theta$

iff $\underline{T}(\underline{x}) = \underline{T}(\underline{y})$, then $\underline{T}(Y)$ is

a minimal sufficient statistic for θ .

Note: want $R(\theta)$ constant as a function of θ iff $\underline{T}(\underline{x}) = \underline{T}(\underline{y})$.

Actually want $f(\underline{x}|\theta) = c_{\underline{x}, \underline{y}} f(\underline{y}|\theta) \quad \forall \theta \in \Theta$

so ignore θ such that $R_{\underline{x}, \underline{y}}(\theta) = \frac{0}{0}$.

But $R_{\underline{x}, \underline{y}}(\theta) = 0$ or $R_{\underline{x}, \underline{y}}(\theta) = \infty$ means \underline{T} is not min suff.

25) ¹¹²⁻³ Know Th 4.5 Let Y_1, \dots, Y_n be iid from 550 29
 a k parameter exp family

$$f(y|\theta) = h(y) c(\theta) \exp[w_1(\theta) t_1(y) + \dots + w_k(\theta) t_k(y)]$$

with natural parameterization

$$f(y|\underline{n}) = h(y) b(\underline{n}) \exp(\eta_1 t_1(y) + \dots + \eta_k t_k(y))$$

(As usual, assume $\Omega = \{\underline{n} \mid \eta_i = w_i(\theta), \theta \in \Theta\}$
 except for the χ^2 and inverse Gaussian distributions.

Let $\underline{T}(\underline{y}) = \left(\sum_{i=1}^n t_1(y_i), \dots, \sum_{i=1}^n t_k(y_i) \right) =$
 $(T_1(\underline{y}), \dots, T_k(\underline{y}))$. Then $\underline{T}(\underline{y})$ is minimal sufficient for
 θ and \underline{n} if the $\eta_i = w_i(\theta)$ do
 not satisfy a linearity constraint.

$\underline{T}(\underline{y})$ is complete sufficient for θ and \underline{n}
 if Ω contains a k dimensional rectangle.

26) Know Cor 4.6 p174] If Y_1, \dots, Y_n are
 iid from a kP -REF, then $\underline{T}(\underline{y})$ is
 sufficient, minimal sufficient and
 complete sufficient for θ and \underline{n} .

Proof min suff and complete suff follow from Th 4.5 since \underline{n} do not satisfy a
 linearity constraint and the k dim open set Ω contains a k dim
 rectangle
 proof of sufficiency $f(\underline{y}|\theta) = \prod_{j=1}^n f(y_j|\theta) =$

$$\underbrace{\left[\prod_{j=1}^n h(y_j) \right]}_{h_3(\underline{y}) \geq 0} \underbrace{\left[c(\theta) \right]^n \exp \left[w_1(\theta) \sum_{j=1}^n t_1(y_j) + \dots + w_k(\theta) \sum_{j=1}^n t_k(y_j) \right]}_{g(\underline{T}(\underline{y})|\theta) \geq 0}$$

So true by factorization.

proof of min suff for Th 4.5

like HW6 problem 6
HW6# 28.5

$$R_{\underline{x}, \underline{y}}(\underline{n}) = \frac{f(\underline{x}|\underline{n})}{f(\underline{y}|\underline{n})} = \frac{\prod_j h(x_j) [\underline{b}(\underline{n})]^n \exp\left[\sum_{i=1}^k n_i T_i(\underline{x})\right]}{\prod_j h(y_j) [\underline{b}(\underline{n})]^n \exp\left[\sum_{i=1}^k n_i T_i(\underline{y})\right]}$$
$$= \frac{\prod_j h(x_j)}{\prod_j h(y_j)} \exp\left[\sum_{i=1}^k n_i (T_i(\underline{x}) - T_i(\underline{y}))\right]$$

is equal to a constant wrt \underline{n} iff

$$\sum_{i=1}^k n_i [T_i(\underline{x}) - T_i(\underline{y})] = \sum_{i=1}^k n_i a_i = c_{\underline{x}, \underline{y}} \forall \underline{n}$$

Since the n_i do not satisfy a linearity constraint, $R_{\underline{x}, \underline{y}}(\underline{n}) = c_{\underline{x}, \underline{y}} \forall \underline{n}$ iff all

$$a_i = 0 \text{ iff } T_i(\underline{x}) = T_i(\underline{y}) \text{ iff}$$

$$\underline{T}(\underline{x}) = \underline{T}(\underline{y}),$$

so $\underline{T}(\underline{Y})$ is minimal sufficient by LSM.

27] Know Tips for finding a minimal sufficient statistic.

- i) Use Factorization to find a candidate suff stat \underline{T} with k as small as possible. Check that \underline{T} is min suff with LSM.
- ii) Recognize or prove that the family is a KP-REF and find $\underline{T}(\underline{Y})$.
- iii) Recognize or prove that the family is

a \mathcal{K} par exp fam (not necessarily REF) 580 29

and a) apply LSM to $T(\underline{y}) = (\sum_j t_1(y_j), \dots, \sum_j t_k(y_j))$

or b) apply Th 4.5 by showing that n_1, \dots, n_k do not satisfy a linearity constraint.

iv) If the support depends on unknown parameter θ , often $Y(1)$ and/or $Y(n)$ will be a component of the min suff stat.

v) If the family is not an exp family and after factorization $k=n$, try LSM on the order statistics $(Y(1), \dots, Y(n))$ (often very hard).

ex) $Y \sim \text{Weibull}(\phi, \lambda)$. min suff stat is $(Y(1), \dots, Y(n))$. But if ϕ is known

$$f(\underline{y} | \lambda) = \prod_{i=1}^n f(y_i) = \prod_{i=1}^n \frac{\phi}{\lambda} y_i^{\phi-1} e^{-y_i^\phi / \lambda} I(y_i > 0)$$

$$= \left(\frac{\phi}{\lambda}\right)^n \prod_{i=1}^n y_i^{\phi-1} e^{-\frac{1}{\lambda} \sum_{i=1}^n y_i^\phi} \prod I(y_i > 0)$$

constant w.r.t λ if ϕ is known

$$\text{So } R_{\underline{x}, \underline{y}}(\lambda) = \frac{f(\underline{x} | \lambda)}{f(\underline{y} | \lambda)} = \frac{\prod I(x_i > 0) \prod x_i^{\phi-1} e^{-\frac{1}{\lambda} (\sum x_i^\phi - \sum y_i^\phi)}}{\prod I(y_i > 0) \prod y_i^{\phi-1} e^{-\frac{1}{\lambda} \sum y_i^\phi}}$$

$\stackrel{\text{set}}{=} C_{\underline{x}, \underline{y}} \lambda$

iff $e^{-\frac{1}{\lambda} (\sum x_i^\phi - \sum y_i^\phi)} \equiv C' \cdot \forall \lambda$ iff

$$\sum x_i^\phi = \sum y_i^\phi. \quad \therefore T(\underline{y}) = \sum_{i=1}^n y_i^\phi \text{ is}$$

min suff by LSM. Note that $f(y | \lambda)$

$$= \underbrace{y^{\phi-1} I(y > 0)}_{h(y)} \underbrace{\frac{\phi}{\lambda}}_{g(\lambda)} \exp\left[\underbrace{-\frac{1}{\lambda} y^\phi}_{w(\lambda) + t(y)} \right]$$

is a IPREF if ϕ is known. 29.5

So $T(\underline{y}) = \sum_{i=1}^n t(Y_i) = \sum_{i=1}^n Y_i \phi$ is min suff and complete suff.

28) In using LSM, it is crucial to treat \underline{x} and \underline{z} as constants. The only variable in $R_{\underline{x}, \underline{z}}(\theta)$ is θ .

Analogy
MLE
 $L(\theta; \underline{x})$

ex] $f(x) = \frac{\log(\theta) \theta^x}{\theta - 1} \quad 0 < x < \infty, \theta > 1$

$$L(\theta) = \frac{\log(\theta)}{\theta - 1} \underbrace{I(x > 0)}_{h(x) \geq 0} \underbrace{\exp(\log(\theta) x)}_{w(\theta) t(x)} \quad n=1$$

$L(\theta) > 0$

$\eta = \log(\theta), \theta > 1 \Rightarrow \Omega = (0, \infty)$

contains 1dim
rectangle = interval

So $\sum_{i=1}^n X_i$ is complete suff

by Th 4.5.

29) Tips for proving $\underline{T}(\underline{y})$ is complete.

i) Note that g can depend on n since n is known, but g can't depend on θ .

ii) If \underline{y} is a kP-REF with $\dim(\Theta) = k$, then $\underline{T}(\underline{y}) = \left(\sum_{j=1}^n t_1(Y_j), \dots, \sum_{j=1}^n t_k(Y_j) \right)$

is complete.

iii) If Y_1, \dots, Y_n are iid $U(\theta_1, \theta_2)$,

For
exam
try
LSM
on $\underline{T}(\underline{y})$