

M582 old Q1

Math 595

Quiz 2 Fall 2005

Name _____

1) Suppose that Y_1, \dots, Y_n are iid with $E(Y) = (1-\rho)/\rho$ and $\text{VAR}(Y) = (1-\rho)/\rho^2$ where $0 < \rho < 1$. a) Find the limiting distribution of $\sqrt{n} \left(\bar{Y}_n - \frac{1-\rho}{\rho} \right)$.

$\xrightarrow{L} N\left(0, \frac{1-\rho}{\rho^2}\right)$ by the CLT.

b) Find the limiting distribution of $\sqrt{n} [g(\bar{Y}_n) - \rho]$ for appropriate function g .

want $g(\mu)$ so that $g\left(\frac{1-\rho}{\rho}\right) = \rho$. Now $\frac{1-\rho}{\rho} = \frac{1}{\rho} - 1$

So $g(\mu) = \frac{1}{1+\mu}$. Note that $g\left(\frac{1-\rho}{\rho}\right) = \frac{1}{1+\frac{1-\rho}{\rho}} = \frac{\rho}{\rho+1-\rho} = \rho$

By the delta method, $\sqrt{n} [g(\bar{Y}) - g(\mu)] \xrightarrow{L} N\left(0, [g'(\mu)]^2 \tau^2\right)$ if $\sqrt{n} (\bar{Y} - \mu) \xrightarrow{L} N(0, \tau^2)$. Here $\mu = \frac{1-\rho}{\rho}$

$g(\mu) = \frac{1}{1+\mu}$, $g'(\mu) = \frac{-1}{(1+\mu)^2}$ and $[g'(\mu)]^2 = \frac{1}{(1+\mu)^4}$

$\text{so } [g'\left(\frac{1-\rho}{\rho}\right)]^2 = \frac{1}{\left(1+\frac{1-\rho}{\rho}\right)^4} = \frac{1}{\left(1+\frac{1}{\rho}-1\right)^4} = \rho^4$

$\text{so } \sqrt{n} [g(\bar{Y}) - \rho] \xrightarrow{L} N\left(0, \frac{\rho^4}{\rho^2} (1-\rho)\right) = N\left(0, \rho^2 (1-\rho)\right)$

$\rightarrow N\left(0, \frac{1-\rho}{\rho^2} \left[\frac{1}{\left(\frac{1-\rho}{\rho}\right)^2}\right]\right)$

1) Let $X_n \sim \text{Binomial}(n, p)$ where the positive integer n is large and $0 < p < 1$.

a) Find the limiting distribution of $\sqrt{n} \left(\frac{X_n}{n} - p \right)$.

$$X_n \stackrel{D}{=} \sum_{i=1}^n Y_i \quad \text{where } Y_i \sim \text{Bin}(1, p), \quad E Y_i = p \quad \text{Var } Y_i = p(1-p)$$

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So by CLT $\sqrt{n} \left(\frac{X_n}{n} - p \right) \xrightarrow{L} N(0, p(1-p)).$

b) Find the limiting distribution of $\sqrt{n} \left[\left(\frac{X_n}{n} \right)^2 - p^2 \right]$.

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Let $g(p) = p^2$. Then $g'(p) = 2p$ and by the delta method $\sqrt{n} \left(\frac{X_n}{n} - p \right) \xrightarrow{L} N(0, p(1-p) \underbrace{(g'(p))^2}_{4p^2})$
 $= N(0, 4p^3(1-p)).$

c) Let $g(\theta) = \theta^3 - \theta$. Find the limiting distribution of $n \left[g\left(\frac{X_n}{n}\right) - c \right]$

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for appropriate constant c when $p = \frac{1}{\sqrt{3}}$. Hint: See (2.5.14).

If $g'(\theta) = 0$, $g''(\theta) \neq 0$ and $\sqrt{n} (T_n - \theta) \xrightarrow{L} N(0, \gamma^2(\theta))$
 then $n [g(T_n) - g(\theta)] \xrightarrow{L} \frac{1}{2} \gamma^2(\theta) g''(\theta) \chi_1^2.$

Now $g'(\theta) = 3\theta^2 - 1$, $g''(\theta) = 6\theta$

$$g\left(\frac{1}{\sqrt{3}}\right) = \left(\frac{1}{\sqrt{3}}\right)^3 - \frac{1}{\sqrt{3}} = \frac{1}{\sqrt{3}} \left(\frac{1}{3} - 1\right) = -\frac{2}{3\sqrt{3}}, \quad g'\left(\frac{1}{\sqrt{3}}\right) = 0, \quad g''\left(\frac{1}{\sqrt{3}}\right) = \frac{6}{\sqrt{3}}$$

$$\gamma^2(p) = p(1-p) \text{ by a). So } \gamma^2\left(\frac{1}{\sqrt{3}}\right) = \frac{1}{\sqrt{3}} \left(1 - \frac{1}{\sqrt{3}}\right).$$

$$\text{So } n \left[g\left(\frac{X_n}{n}\right) - \underbrace{\left(-\frac{2}{3\sqrt{3}}\right)}_c \right] \xrightarrow{L} \frac{1}{2} \frac{1}{\sqrt{3}} \left(1 - \frac{1}{\sqrt{3}}\right) \frac{6}{\sqrt{3}} \chi_1^2$$

$$= \left(1 - \frac{1}{\sqrt{3}}\right) \chi_1^2$$

$$\frac{1}{2} \gamma^2(p) g''(p) = \frac{1}{2} p(1-p) 6p = 3p^2(1-p) = 3\left(\frac{1}{\sqrt{3}}\right)^2 \left(1 - \frac{1}{\sqrt{3}}\right) = 1 - \frac{1}{\sqrt{3}}$$