

such that

$$n^\delta (T_{1n} - \theta) \xrightarrow{D} N(0, \sigma_1^2(F))$$

$$\text{and } n^\delta (T_{2n} - \theta) \xrightarrow{D} N(0, \sigma_2^2(F)).$$

Then the asymptotic relative efficiency  
of  $T_{1n}$  wrt  $T_{2n}$  is

$$\text{ARE}(T_{1n}, T_{2n}) = \frac{\sigma_2^2(F)}{\sigma_1^2(F)}.$$

53] a)  $0 < \delta \leq 1$  and usually  $\delta = \frac{1}{2}$ .

b) If  $T_{1n}$  has  $n^{\delta_1}$ , then

$T_{1n}$  is better than  $T_{2n}$  if  $\delta_1 > \delta_2$ .

c) Often 2 estimators  $T_{1n}$  and  $T_{2n}$   
do not estimate the same parameter  
 $\theta$  for many distributions  $F$ , unless

The distributions are symmetric about  $\theta$ . 23.5

c) For fixed  $F$ )

$T_{1n}$  is better than  $T_{2n}$  if  $ARE \geq 1$   
worse  $ARE < 1$ .

So the better estimator has  
the smaller  $\sigma_i^2(F)$ .

ex) Let  $Y_1, \dots, Y_n$  be iid  $N(\mu, \sigma^2)$ .  
Let  $\theta = \mu = E(Y) = MED(Y)$ .

Let  $T_{1n} = \bar{Y}_n$  and  $T_{2n} = MED(n)$ ,

By the CLT,  $\sigma_1^2(F) = \sigma^2$ .

By 49),  $\sigma_2^2(F) = \frac{1}{4[\bar{F}(MED(Y))]^2}$

$$= \frac{4 \left[ \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left(\frac{-\theta}{2\sigma^2}\right) \right]^2}{1 - \exp\left(-\frac{\theta}{\sigma^2}\right)}$$

$$= \frac{2\pi\sigma^2}{4} = \frac{\pi\sigma^2}{2}$$

$$\text{ARE}(T_{1n}, T_{2n}) = \text{ARE}(T_n, \text{MED}(n))$$

$$= \frac{(\pi\sigma^2/2)}{\sigma^2} = \frac{\pi}{2} \approx 1.5708$$

Hence  $\bar{T}_n$  is a better estimator of  $\mu$  than  $\text{MED}(n)$  for the family of  $N(\mu, \sigma^2)$  distributions,  $\sigma^2 > 0$ .

54) p485-6 know Assume  $T'(\theta) \neq 0$ , 245

An estimator  $T_n$  of  $T(\theta)$  is

asymptotically efficient if

$$\sqrt{n} (T_n - T(\theta)) \xrightarrow{D} N\left(0, \frac{\mathbb{E}[T(\theta)]^2}{I_1(\theta)}\right)$$

$$\sim N(0, \text{FCRLB}_1(T(\theta))).$$

In particular, an estimator

$T_n$  of  $\theta$  is asymptotically efficient

$$\text{if } \sqrt{n}(T_n - \theta) \xrightarrow{D} N\left(0, \frac{1}{I_1(\theta)}\right) \sim N\left(0, \text{FCRLB}_1(\theta)\right).$$

55) p486 If  $T_{2n}$  is an

asymptotically efficient estimator of  $\theta$ ,

if  $I_1(\theta)$  and  $N(\theta)$  are continuous

functions, and if  $\sqrt{n}(T_{2n} - \theta) \xrightarrow{D} N(0, N(\theta))$ ,

Then under regularity

conditions  $N(\theta) \geq \frac{1}{I_1(\theta)}$  and

$$\text{so } \text{ARE}(T_{1n}, T_{2n}) = \frac{\left(\frac{1}{I_1(\theta)}\right)}{N(\theta)} =$$

$\frac{1}{I_1(\theta) N(\theta)} \leq 1$ . Hence asymptotically efficient estimators are better than estimators of the form  $T_{1n}$ .

56) The standard limit theorem (theorem 40) says that under strong regularity conditions a) The MLE  $\hat{\theta}_n$  of  $\theta$  and the MLE  $\hat{T}(\hat{\theta}_n)$  of  $T(\theta)$  are asymptotically efficient estimators of  $\theta$  and  $T(\theta)$ .  
b)  $\hat{T}(\hat{\theta}_n)$  is an asymptotically

efficient estimator of  $T(\theta)$

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if  $\hat{\theta}_n$  is the UMVUE of  $\theta$ ,

57} P70 Th If  $n^s(T_n - T(\theta)) \xrightarrow{D} X_\theta \quad \forall \theta \in \Theta$

then  $T_n$  is a consistent estimator of  $T(\theta)$

where  $0 < s \leq 1$ ,

Note:  $X_\theta \sim N(0, \sigma(\theta))$  and  $s=0.5$  are common.

If  $\sqrt{n}(T_n - \theta) \xrightarrow{D} X_\theta$ , then  $T_n \xrightarrow{P} \theta$ .

Know P70

58} Slutsky's Theorem: Suppose

$y_n \xrightarrow{D} Y$  and  $w_n \xrightarrow{P} c$  for  
some constant  $c$ . Then

a)  $y_n + w_n \xrightarrow{D} Y + c$

b)  $y_n w_n \xrightarrow{D} c Y$

c)  $\frac{y_n}{w_n} \xrightarrow{D} \frac{Y}{c}$  if  $c \neq 0$ .

Note:  $w_n \xrightarrow{P} c$  iff  $w_n \xrightarrow{D} c$ ,

regular convergence

If  $c_n \rightarrow c$  then  $c_n \xrightarrow{\text{wpl}} c$  and  $c_n \xrightarrow{P} c$ .

If  $w_n \xrightarrow{A} c$ ,  $A = \text{wpl or } R$ , then  
 $w_n \xrightarrow{P} c$ .

Ex) If  $y_1, \dots, y_n$  are iid with

$E(y_i) = \mu$  and  $V(y_i) = \sigma^2$   
then  $s_n^2 \xrightarrow{P} \sigma^2$  and  $s_m^2 \xrightarrow{P} \sigma^2$ .

ex) a)  $s_n^2 \xrightarrow{P} \sigma^2 \therefore s_m^2 = \frac{n-1}{n} s_n^2 \xrightarrow{w} 1(\sigma^2) = \sigma^2$

by Slutsky's Th.

b)  $s_m^2 \xrightarrow{P} \sigma^2 \therefore s_n^2 = \frac{n-1}{n} s_m^2 \xrightarrow{P} 1(\sigma^2) = \sigma^2$

c) If  $y_n \xrightarrow{A} d$  and  $w_n \xrightarrow{B} c$  where  
 $c$  and  $d$  are constants and

$A$  or  $B$  are wpl,  $R$ ,  $P$  or  $D$ , then

$y_n \xrightarrow{D} d$  and  $w_n \xrightarrow{P} c$  so

$$Y_n + W_n \xrightarrow{P} d + C$$

$$Y_n W_n \xrightarrow{P} dC$$

$$\frac{Y_n}{W_n} \xrightarrow{P} \frac{d}{C} \text{ if } C \neq 0.$$

In particular, if  $Y_n \xrightarrow{P} d$  and  $W_n \xrightarrow{P} C$   
 the above 3 results hold, and P  
 can be replaced by D.

d) If  $Y_n \sim t_n$ , a t dist with n  
 degrees of freedom then  $t_n \xrightarrow{D} Z \sim N(0,1)$

Proof:  $Y_n \stackrel{D}{=} \frac{Z}{\sqrt{n}/t_n}$  where  $Z \perp\!\!\!\perp V_n \sim \chi^2_n$ .

$V_n \stackrel{D}{=} \frac{\sum_{i=1}^n X_i}{n}$  where  $X_i \stackrel{iid}{\sim} \chi^2_1$ .

So  $\frac{V_n}{n} \xrightarrow{P} 1 \quad \therefore \sqrt{\frac{V_n}{n}} \xrightarrow{P} \sqrt{1} = 1$  by 583

e) If  $Y_n \sim F_{p, d_n}$ , an F dist with p  
 and  $d_n$  degrees of freedom then

$PY_n = P F_{pdn} \xrightarrow{D} X_p^2$  if  $d_n \rightarrow \infty$  as  $n \rightarrow \infty$ .

Proof  $Y_n \sim F_{pdn} \stackrel{D}{=} \frac{X_1/p}{X_2/d_n}$  where

$$X_1 \sim X_p^2 \quad \text{if } X_2 \sim X_{dn}^2$$

e.g.  $X_2 \stackrel{D}{=} \sum_{i=1}^{d_n} Y_i$ ,  $Y_i \stackrel{iid}{\sim} X_1^2$  and

$$\frac{X_2}{d_n} \xrightarrow{P} 1 \quad \therefore PY_n = \frac{1}{\frac{X_2}{d_n}} X_1 \xrightarrow{D} 0 \quad X_1 \stackrel{D}{=} X_p^2$$

constant

58) Let  $g: \mathbb{R} \rightarrow \mathbb{R}$  be continuous at  $c$ .

a) If  $x_n \xrightarrow{P} c$ , then  $g(x_n) \xrightarrow{D} g(c)$

b) If  $x_n \xrightarrow{P} c$ , then  $g(x_n) \xrightarrow{P} g(c)$

c) If  $x_n \xrightarrow{wpl} c$ , then  $g(x_n) \xrightarrow{wpl} g(c)$ .

Note: If  $x_n \xrightarrow{P} c$ , then  $x_n \xrightarrow{P} c$  so

$g(x_n) \xrightarrow{P} g(c)$ .

$$Ex: x_n \xrightarrow{P} \mu \Rightarrow E(x_n^2) \xrightarrow{P} E(\mu^2)$$

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59) Generalized Continuous Mapping Th

If  $x_n \xrightarrow{D} X$  and the function  $g$

is such that  $P[\bar{x} \in C(g)] = 1$

where  $C(g)$  is the set of points  
at which  $g$  is continuous,  
then  $g(x_n) \xrightarrow{D} g(X)$ .

Note:  $P[\bar{x} \in C(g)] = 1 \Leftrightarrow P[\bar{x} \in D(g)] = 0$

where  $D(g)$  is the set of points where  
 $g$  is not continuous.

ex) If  $x_n \xrightarrow{D} X$  then  $\frac{1}{x_n} \xrightarrow{D} \frac{1}{X}$   
continuous RV

If  $X$  has a pdf since  $P(X=0) = 0$

and  $X=0$  is the only discontinuity  
point of  $g(x) = \frac{1}{x}$ .

$\frac{1}{x_n} \xrightarrow{D} \frac{1}{x}$ , if  $x$  has a pmt  
with  $P(x=0) = 0$ .

### 603 Continuous mapping Th

If  $x_n \xrightarrow{D} x$  and  $g$  is a continuous function, then  $g(x_n) \xrightarrow{D} g(x)$ .

Note} 603 is a special case of 59  
where  $D(g) = \emptyset$ .

613 A complex RV  $Z = X + iY$   
where  $X$  and  $Y$  are ordinary RVs.

Then  $E(Z) = E(X) + iE(Y)$  where  
 $i = \sqrt{-1}$ . The modulus

$|Z| = \sqrt{X^2 + Y^2}$ . Much of the  
theory for  $Z$  parallels that of  $X$ .

In particular, (linearity)

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$$|E(z)| \leq E|\bar{z}|, \quad z_1 \neq z_2$$

$\Rightarrow h_1(z_1) \perp\!\!\!\perp h_2(z_2)$  where  $h_i$  is a function of  $z_i$  alone.

Often  $z = e^{itX} = \cos(tX) + i \sin(tX)$ .

62) know The characteristic function

$$C_X(t) = E(e^{itX}) = E[\cos(tx)] + iE[\sin(tx)]$$

real RV

63)  $C_X(t)$  always exists and completely determines the dist of  $X$ .

i)  $C(0) = 1$

ii)  $|C(t)| \leq 1$  where the modulus

$$|a+ib| = \sqrt{a^2+b^2}$$

iii)  $C(t)$  is a continuous function.