

25) take the sample mean of each variable and collect in a vector.

know for ED			bisoral	see H(13)
length	Nasal ht			
X <sub>1</sub>	X <sub>2</sub>	X <sub>3</sub>		
168	51	104	1st person	
176	55	106		n = 5
179	55	104		
176	46	114		
177	48	101	5th person	

$\sum X_{1i} = 876$     $\sum X_{2i} = 255$     $\sum X_{3i} = 529$

$$\bar{X} = \frac{1}{5} \begin{pmatrix} 876 \\ 255 \\ 529 \end{pmatrix} = \begin{pmatrix} 175.2 \\ 51.0 \\ 105.8 \end{pmatrix}$$

sample mean of length

know for ED  
ex) Let T be the coordinate wise median, MED( $\mathbb{R}^2$ ).

$$T = \begin{pmatrix} 176 \\ 51 \\ 104 \end{pmatrix} \quad \begin{matrix} 168 & 176 & 176 & 177 & 179 \\ 46 & 48 & 51 & 55 & 55 \\ 101 & 104 & 104 & 106 & 114 \end{matrix}$$

26) The R function cov.mcd gives estimator  $(\hat{\mu}, \hat{\Sigma})$  that is often useful when outliers are present, but RCF and RMVN are faster and backed by theory.

§ 13.6

21D

plot

27) The  $j$ -th start  $(T_{-1,j}, C_{-1,j})$  is an initial MLE estimator. Then  $(T_{0,j}, C_{0,j}) = (\bar{x}_{0,j}, s_{0,j})$  is the classical estimator computed from the  $C_n \approx \frac{1}{2}$  cases with the smallest  $D_i$   $(T_{-1,j}, C_{-1,j})$ . Repeat the iteration for  $t$  steps resulting in the sequence of estimators  $(T_{-1,j}, C_{-1,j}) (T_{0,j}, C_{0,j}), \dots, (T_{t,j}, C_{t,j}) \rightarrow (T_{\infty,j}, C_{\infty,j}) = (\bar{x}_{k,j}, s_{k,j})$  the  $j$ -th attractor. The concentration estimator  $(T_A, C_A)$  is the attractor used to obtain the final estimators  $j=1, \dots, E$ .

28) PROOF  $\det(C_{t+1,j}) \leq \det(C_{t,j})$  with equality iff  $(T_{t,j}, C_{t,j}) = (T_{t+1,j}, C_{t+1,j})$  for  $t \geq 0$ . So the determinant decreases until convergence.

29) PROOF The volume of the hyperellipsoid

$$\left\{ \mathbf{z} : (\mathbf{z} - \bar{x}_{k,j})^T S_{k,j}^{-1} (\mathbf{z} - \bar{x}_{k,j}) \leq h^2 \right\} =$$

$$2\pi^{\frac{p}{2}} h^p \sqrt{\det(S_{k,j})}. \text{ So small volume goes with}$$

$\underbrace{P(\frac{p}{2})}_{\text{constant}}$   
small determinant.

30) For EC distributions, the regions of highest density are  $\left\{ \mathbf{z} : (\mathbf{z} - \bar{x})^T S^{-1} (\mathbf{z} - \bar{x}) \leq D_{\text{for}}(x, \frac{1}{2}) \right\}$  are hyperellipsoids

of  $\hat{\pi}$ . If  $(\hat{\pi}, \hat{c})$  is a consistent estimator  
of  $(\pi_0, s^2)$  for some  $s > 0$ , then

$\hat{\pi}_i : D_i^2(\hat{\pi}, \hat{c}) \leq \text{MED}(D_i^2(\pi_0, c)) \}$  estimates the  
highest 50% density region.

31) Using the concentrated attractor

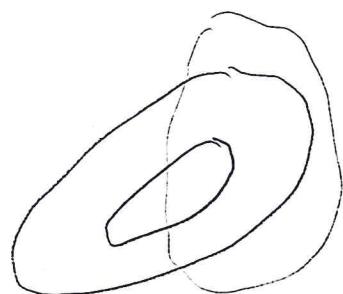
iterated to convergence seems

to estimate a highest

density region for EC data + outliers

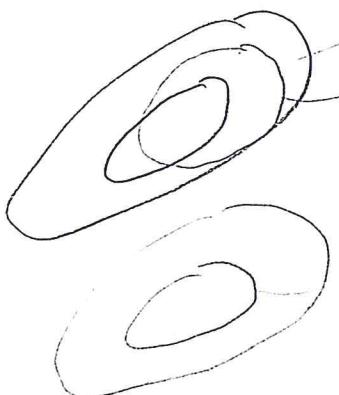
if the attractor region does not contain any outliers,

(If there are 40% outliers, then the highest  
density region containing  $\frac{2}{3}$  of the clean data is estimated)  
 $\frac{2}{3}60\% = 50\%$



Step 0 Scale ellipsoid to contain half the  
data take classical estimator

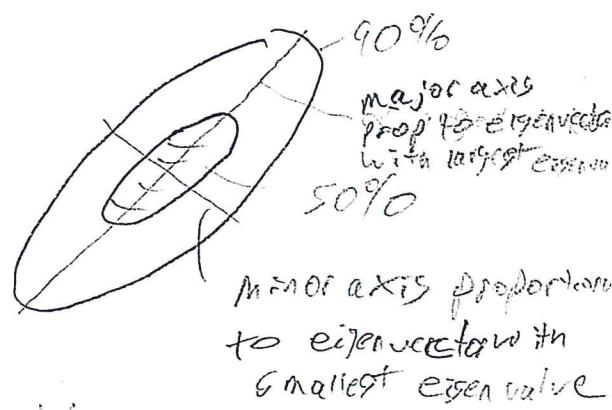
cover half the data



Step 1 new region is a lot better

at convergence, region for attractor seems  
pretty good

32)  $\hat{\pi}$  determines the shape of the highest  
density region which is a hyper-ellipsoid  
for EC distributions (with  $g \downarrow$ ),



w) an elemental start for MLD uses

$p+1$  randomly selected cases  $T_j$ . Then

$(T_{j*}, C_{j*}) = (\bar{X}_{j*}, S_{j*})$  = classical estimator applied to the  $p+1$  cases.  $\text{COV}_{\text{MCD}} = F_{\text{MCD}}$  uses 500 randomly chosen elemental starts.

For each of the  $K=500$  attractors, find  $\det(C_{j,k}), \dots, \det(C_{500,k})$ . Suppose  $j=m$  corresponds to the attractor with the minimum determinant.

Then  $(T_{\text{MCD}}, C_{\text{MCD}})$  uses the attractor  $(T_{m*}, C_{m*})$ .

BB) The DGT<sub>t</sub> estimator is very simple.

Let  $(\bar{X}_1, C_1) = (\bar{X}, S) =$  classical estimator

be the start. Then  $(T_{\text{DGT}}, C_{\text{DGT}})$  is the attractor.

Using  $k=10$  concentration steps works well. The DGT<sub>t</sub> estimator has much more outlier resistance than the classical estimator.

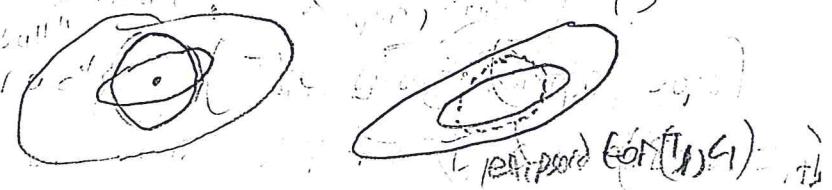
BB) PIDE The MB (Median Ball) estimator

$(\bar{X}_1, C_1) = (\text{MED}(W), \Sigma)$  as the start where  $\text{MED}(W)$

$\rightarrow$  the coordinatewise median. Hence  $(\bar{x}_0, \bar{\sigma}_0)$  is the classical estimator applied to the  $n \times \frac{n}{2}$  cases closest to  $\text{MED}(W)$  in Euclidean distance.

Use  $K=3$  concentration steps. This estimator is

a highly outlier resistant estimator. It can also be obtained with the



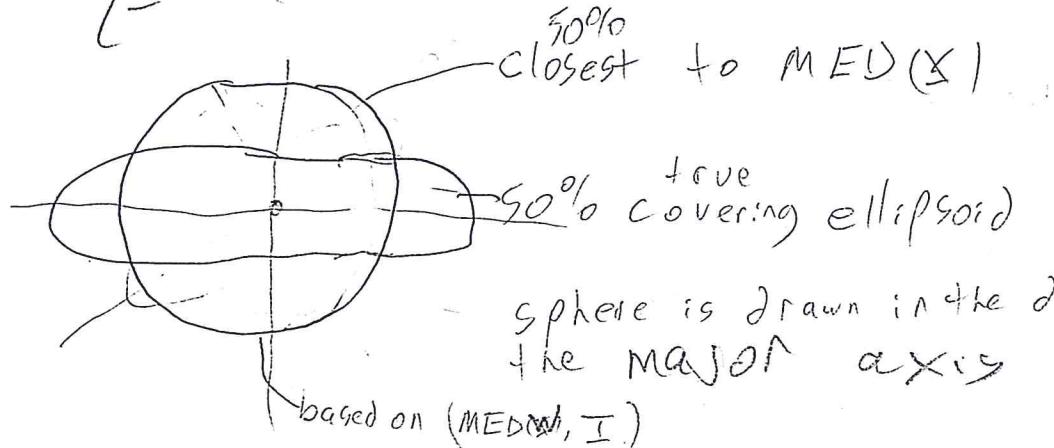
38) The FCH estimator uses the misattraction factor if  $T_{DBK}$  has a greater Euclidean distance from MED(w) than half the data (so  $T_{DBK}$  is outside  $\text{MED}(w)$ ) than half the data (so  $T_{DBK}$  is outside one of the median ball that contains half of the data). Otherwise FCH uses the DBK or MB estimator that has the smallest determinant  $(\bar{\tau}_A, c_A)$ .

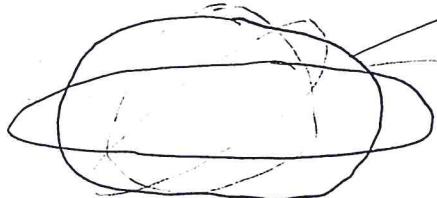
$$\text{Then } \bar{\tau}_{FCH} = \bar{\tau}_{DBK} \text{ and } C_{FCH} = \frac{\text{MED}(D^2(\bar{\tau}_A, c_A))}{\chi^2_{P,15}} c_A.$$

$D_{DBK} < D_{MB} \Rightarrow \text{det } D_{DBK} < \text{det } D_{MB}$   
otherwise MB

37) On a large class of EC distributions, the prob that  $(\bar{\tau}_A, c_A) = (\bar{\tau}_{DBK}, c_{DBK})$  goes to one as  $n \rightarrow \infty$ , and  $(\bar{\tau}_{FCH}, c_{FCH})$  is a  $\sqrt{n}$  consistent estimator of  $(\bar{\mu}, c^*)$  where  $c > 0$  and  $c^* = 1$  for MVN data.

38) The median ball attractor results in ellipsoids  $\{x | (x - \bar{\tau}_{MB})^\top E_{MB}^{-1} (x - \bar{\tau}_{MB}) \leq d^2\}$  that are "too short" in the major axis and "too fat" in the minor axis compared to  $\{x | (x - \bar{\mu})^\top (x - \bar{\mu}) \leq d^2\}$ .

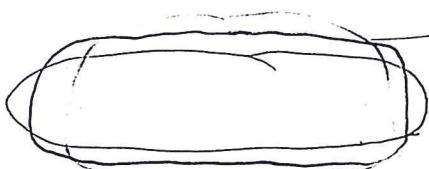




(DGP, 40)

29

too fat on the minor axis  
too short on the major axis



based on

$$(\hat{T}_{MB}(x)) \neq (\hat{T}_{MB}(y))$$

better but still biased, if

(PC concentration steps are used)

27 (It is not known whether MB is biased or not) if concentration  
is iterated to convergence)

For the next section, we will assume no bias.

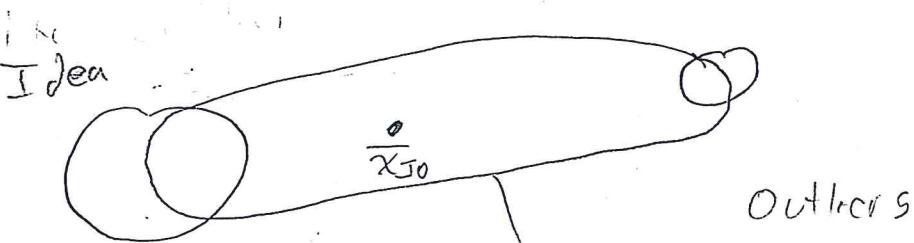
The DGT estimator is not outlier resistant.

By 19, the MAA estimator is consistent.

39 When outliers are present "that", the DGT estimator can't detect, usually  $\det(C_{MB}) < \det(C_{MAA})$  so the FCA estimator is highly outlier resistant.

40 Even if the 50% of cases with the smallest distances based on the start contains outliers, the attractor may use a half set without outliers.

DGP



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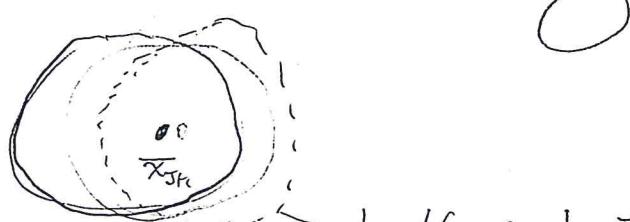
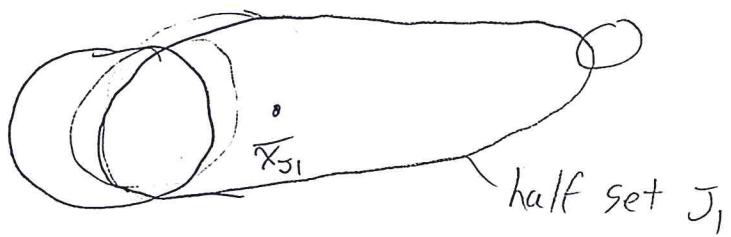
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PCL XL error

If the number of outliers in the half set  $J < 25\%$ , the  $\bar{x}_J$  is closer to the bulk.

at the outliers than to the outliers. After a concentration step, fewer outliers and more clean cases will be used (if the outliers are far enough away). After k concentration steps, the half set may be clean.



clean data "won the tug-of-war". HW7C illustrates this phenomenon,

half set  $J_k$  contains no outliers.

(ii) The MEFH estimator can give good results even when nearly 50% of the cases are outliers.

(iii) A scatterplot matrix of the distances

from COV.MCD, FCH, DGH and the median ball estimator can be useful. HW7B makes DD plots based on these 4 estimators.

(iv) PVB and RFCH are reweighted inversions of FCH. They are not consistent estimators of  $(\mu, \sigma; \#)$  where  $\sigma = 1$ . For MVN data,  $\hat{\mu}_{PVB} = \hat{\mu}_{RFCH} = \hat{\mu}_{FCH}$  (but  $\hat{\sigma}_{PVB}^2 \neq \hat{\sigma}_{RFCH}^2 = \hat{\sigma}_{FCH}^2$ ). (Also,  $\hat{\sigma}_{PVB}^2$  does not contain the original  $\sigma^2$ ).